# AN OPTIMUM PITCH PROCESSING MODEL FOR SIMULTANEOUS COMPLEX TONES 

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## ABSTRACT

An extension of Goldstein's Optimum Processing Theory is pre- \& sented which can account for pitch perception behavior for simultaneous complex tones. The essence of the theory is that all aurally resolved stimulus frequencies are transformed into independent Gaussian random variables with a variance that depends only on the frequency of each partial. A central processor is assumed to use its prior knowledge about the number of simultaneously present tone complexes and the proper parsing of the observed random variables to find the respective fundamentals of the best fitting harmonic templates. In a series of pitch identification experiments for two simultaneous two-tone complexes with diotically and dichotically distributed partials, some model assumptions and their consequences were tested. It was found that (1) the processes of estimating two simultaneous (missing) fundamentals are to a large extent independent, (2) that the central processor tends to group the partial percepts on the basis of common fundamental and not on the basis of ear input, and (3) that pitch identification performance degrades only noticeably if none of the stimulus partials of both tone complexes are aurally resolved.

## INTRODUCTION

The problem how we perceive the pitch of complex tones has kept psychoacousticians busy for more than a century. In particular the problem of the so called "missing fundamental", a pitch percept that corresponds with the fundamental frequency of a harmonic tone complex while that complex actually has only overtones, has been the object of many experimental and theoretical studies. Various pieces of important empirical evidence and theories to account for such evidence have been brought forward by Seebeck [1], Ohm [2], Helmholtz [3], Fletcher [4], Schouten [5] and Békésy [6].
More recent experiments by Plomp [7], Ritsma [8] and Houtsma and Goldstein [9] have progressively shown that the real cause of the "missing fundamental" phenomenon must not be sought in the peripheral, but rather in the central part of the auditory system. The new experimental evidence has led to the formulation of some new central pitch theories, of which the Virtual Pitch Theory of Terhardt [10] and the Optimum Processor Theory of Goldstein [11] are the principal variants. These theories were developed and quantified mostly on the basis of pitch perception data obtained with isolated complex tones or short sequences of
such tones.

In music, especially in the Western hemisphere, we usually deal with harmonic or polyphonic sound patterns in which either a melody is accompanied with chords or several melodies are played simultaneously against one another. This poses the interesting problem how our auditory system is able to perceive two or more simultaneous pitches when it is acoustically exposed to a cluster of harmonics that belong to several different tone complexes. The same problem actually occurs when one tries to track the prosodic contours of two simultaneously spoken sentences or, more realistically, when one tries to follow the pitch contour of one spoken sentence against a background of other speech. Although both central pitch theories mentioned $[10,11]$ are in principle able to cope with this problem, this has never been worked out specifically or tested against systematic empirical data.
In this study the Optimum Processor Theory of Goldstein will be extended and tested with experimentally obtained pitch identification data for two simultaneous complex tones. The model extension will be treated in Sect. I. Descriptions of the experimental procedure and the results are given in Sects. II and III. Computer simulations of model performance are discussed in Sect. IV, and conclusions of the study are presented in Sect. V.

## I. EXTENSION OF TḤE OPTIMUM PROCESSOR THEORY

In Goldstein's Optimum Processor Theory [11] and in a later extension of that theory [12] it was assumed that:

1. the complex tone input in both ears is spectrally analyzed and only frequency information of sufficiently resolved partials is kept; phase and amplitude information is discarded;
2. independent Gaussian random variables $r_{i}$, of zero mean and with variance depending on frequency only, are added to each resolved frequency to form the noisy frequeny codes $x_{i}=f_{i}+r_{i}$; 3. a central processor rank-orders all noisy frequency codes from both ears and performs a maximum-likelihood estimate of the best-fitting harmonic numbers and fundamental of some underlying harmonic complex-tone template.
This model, which was originally formulated to describe perception of a single pitch from a single complex tone, can easily be extended to accomodate identification tasks of pitches from simultaneously sounding complex tones. In this study we will focus on the task of identifying two fundamental pitches in an acoustic stimulus that comprises two simultaneous two-tone complexes, each one having successive harmonics. Extension of the model to other cases, e.g., three or four simultaneous two-tone complexes or two simultaneous multi-tone complexes, is, in principle, not different but may be computationally more complex.

Suppose now that the acoustic stimulus consists of four frequencies: $f_{1}=m f_{01}, f_{2}=(m+1) f_{01}, f_{3}=n f_{02}$, and $f_{4}=(n+$ the auditory system. The frequencies $f_{1}$ through $f_{4}$ are the transformed into four independent Gaussian random variables $x$ drough $x_{4}$, having means of $f_{1}$ through $f_{4}$ respectively, and stanas $\sigma_{i}$, the likelihood function to be optimized by the processor is given by the expression:

$$
\begin{aligned}
L\left(f_{1}, f_{2}, f_{3}, f_{4}\right)= & \frac{1}{4 \pi^{2} \sigma_{1} \sigma_{2} \sigma_{3} \sigma_{4}} \cdot \exp \left[-\frac{\left(x_{1}-f_{1}\right)^{2}}{2 \sigma_{1}^{2}}\right] . \quad(1) \\
& \exp \left[-\frac{\left(x_{2}-f_{2}\right)^{2}}{2 \sigma_{2}^{2}}\right] \cdot \exp \left[-\frac{\left(x_{3}-f_{3}\right)^{2}}{2 \sigma_{3}^{2}}\right] . \\
& \operatorname{expl}\left[\frac{\left(x_{4}-f_{4}\right)^{2}}{\left.2 \sigma_{4}^{2}\right]}\right] .
\end{aligned}
$$

Maximizing Eq.(1) is equivalent to maximizing the log-likelihoo function

$$
\begin{align*}
\Lambda\left(f_{1}, f_{2}, f_{3}, f_{4}\right)= & -\frac{\left(x_{1}-f_{1}\right)^{2}}{\sigma_{1}^{2}}-\frac{\left(x_{2}-f_{2}\right)^{2}}{\sigma_{2}^{2}}  \tag{2}\\
& -\frac{\left(x_{3}-f_{3}\right)^{2}}{\sigma_{3}^{2}}-\frac{\left(x_{4}-f_{4}\right)^{2}}{\sigma_{4}^{2}} .
\end{align*}
$$

In interpreting this log-likelihood function, the knowledge the pro cessor has about the make-up of the stimulus and the task to (A) in which the processor has full knowledge of the fact that there are two two-tone complexes, and hence two pitches to be found, as well as knowledge of the correct parsing, i.e., of the correct harmonic interpretation of each observed input $x_{i}$. We nes present is known, but the correct parsin is unk ormpex processor
Case A . When the number of fundamental pitches to be identified and also all parsing information is available to the processor, it makes the following substitutions in Eq.(2)

$$
\begin{aligned}
& f_{1}=\hat{m} \hat{f}_{01} \\
& f_{2}=(\hat{m}+1) \hat{f}_{01} \\
& f_{3}=\hat{n} \hat{f}_{02} \\
& f_{4}=(\hat{n}+1) \hat{f}_{02}
\end{aligned}
$$

and maximizes the expression with respect to the (lower) harmonic number estimates $\hat{m}$ and $\hat{n}$ and the fundamental pitch estimates $f_{01}$ and $\hat{f}_{02}$. Because of the statistical independer of the input variables $x_{i}$, the first two terms and the last two lerms of Eq. (2) can be maximized separately. The two indepen dent fitting procedures, each one identical to the one described by Goldstein [11], yield the optimum harmonic-number estimates $\dot{m}$ and $\hat{n}$ as well as the fundamental pitch estimates $f_{01}$ and $f_{0}$,
The probabilities $\operatorname{Pr}[\hat{m}=k]$ and $\operatorname{Pr}[\hat{n}=l]$ with $k$ and $l$ bein integers, are discrete probabilities which can be computed from the stimulus frequencies $f_{i}$ and the fixed and known frequency coding noise function $\sigma\left(f_{i}\right)$, and the fundamental pitch estimate

$$
\begin{align*}
& \hat{f}_{01}=\frac{\left[x_{1} / \hat{m}\right]^{2}+\left[x_{2} /(\hat{m}+1)\right]^{2}}{x_{1} / \tilde{\tilde{c}}+x_{2} / /(\hat{m}+1)}  \tag{4}\\
& \hat{f}_{02}=\frac{\left.\left[x_{3} / \hat{n}\right]^{2}+\mid x_{4} /(\hat{n}+1)\right]^{2}}{x_{/ 2} / \hat{n}+x_{4} /(\hat{n}+1)} .
\end{align*}
$$

each trial by pressing two out of five buttons on a response box in any temporal order. There was unlimited response time, and fixed delay.
Four stimulus conditions were investigated. In condition 1, which was diotic, all four stimulus frequencies $f_{1}$ through $f_{4}$ were presented to both ears. In condition 2 , which was dichotic, one note (comprising the frequencies $f_{1}$ and $f_{2}$ ) was presented to one ear, while the other note (with the frequencies $f_{3}$ and $f_{4}$ ) went to $t$, other ear. In condicions 3 and 4, which wo dichotic, dition 3 , one ear received $f_{1}$ and $f_{3}$, while the other ear received $f_{2}$ and $f_{4}$. In condition 4, one ear received $f_{1}$ and $f_{4}$ while the other ear received $f_{2}$ and $f_{3}$.
Since there are ten different combinations of two notes in a total set of five, and since there were $9 \times 9=81$ different harmonic 810 physically different stimuli and a total of 10 different response categories. Each of these stimuli was, on the average, presented six times to each of four subjects, for a total of 4500 identification trials per subject for each stimulus condition.
III. RESULTS

The raw data of all experiments consisted of a record for each trial of the presented fundamentals $f_{01}$ and $f_{02}$, the lower harmonic numbers $m$ and $n$, and the subject's two responses $R_{a}$ and $R_{b}$. A response ( $R_{a}, R_{b}$ ) could be an identification of the perceived fundamentals ( $f_{01}, f_{02}$ ) or ( $f_{02}, f_{01}$ ), since the order of pressing the response buttons was arbitrary
To obtain some insight in the perceptual independence of the idenlification processes for each of the two simultaneous notes, the raw
data were processed by two different methods. In the first method all trials were counted for every ( $m, n$ ) combination where both $f_{01}$ and $f_{02}$ were identified correctly. The results of this way of counting yielded half-matrices of 'percent correct' scores, $\operatorname{Pc}(k, l)$, lor each subject, in which $k$ and $l$ are integers representing the harmonic numbers ( $m, n$ ) or $(n, m$ ). They are hall-matrices bemakes both halves of the matrix mirror images. In the second method only the correct identification of one of the two simultaeous notes was considered as a function of both (lower) harmonic umbers but regardess of the identification response to the other note. The resulting score, designated as $\operatorname{Pc}(k \mid l)$, represents the rcentage correct identifications of fol for $k=m$ and $t=n$, as Was the correct identifications of for for $k=n$ and
Both processed data matrices $\operatorname{Pc}(k, l)$ and $\operatorname{Pc}(k \mid l)$ can be used to Ind an underlying processor performance function $\operatorname{Pr}[\hat{k}=k]$, the processor's probability of correctly estimating the harmonic orde f any complex tone. This was done with a minimum chi-square thing procedure which looked for those $\operatorname{Pr}[k=k]$ functions tha matrices $\mathrm{Pc}(k, l)$ and $\mathrm{Pc}(k l \mid)$. The details of this procedure, which also involved some assumptions about the decision process for the particular experimental paradigm that was used, are discussed in recent publication by the authors [13]:
he functions $\operatorname{Pr}_{1}[\hat{k}=k]$ derived from the matrix $\operatorname{Pc}(k \mid l)$ and $\mathrm{r}_{2}[\hat{K}=k]$ derived from $\mathrm{Pc}(k, l)$ are shown in Fig. 1a-d as tri through 4 squares respectively for the experimental $\hat{\text { condic }}$,
values of $k$ and $\left.\operatorname{Pr}_{1} \mid \hat{\mathrm{K}}=\mathrm{k}\right]<\operatorname{Pr}_{2}[\hat{\mathrm{~K}}=\mathrm{k}]$ for large values of $k$, the pendent in the sense that the perception of the more salient pitch, i.e., the one represented by the lowest harmonic numbers, inhibit correct perception of the less salient pitch [13]. Figure 1a-d show that in condition 2 only subject JH noticeably exhibits this effec of mutual dependence of the two identification processes, but in conditions 1,3 and 4 all subjects except MZ seem to show a smal amount of mutual dependence.


Fig. 1a. The processor's probability of identifying the correct har monic order of a complex tone. The harmonic order is shown on the abscissa. Triangles designate $\operatorname{Pr}_{1}[\hat{\mathrm{k}}=\mathrm{k}]$, squares $\operatorname{Pr}_{2}[\hat{\mathrm{k}}=\mathrm{k}]$

#  

Fig. 1b. Same as Fig la, but computed from data of condition 2 .


Fig. 1c. Same as Fig la, but computed from data of condition 3.


Fig. 1d. Same as Fig la, but computed from data of condition 4.

From either probability function $\left.\operatorname{Pr}_{1} \mid \hat{\mathrm{k}}=\mathrm{k}\right]$ or $\mathrm{Pr}_{2}[\hat{\mathrm{~K}}=\mathrm{k}]$ one can now compute the model's variance function $\sigma(f)$ which represents the frequency coding noise and is its only free parameter. A set of those sigma functions is shown in Fig. 2. The functions were computed from the averaged $\mathrm{Pr}_{1}[\hat{\mathrm{~K}}=\mathrm{k}]$ and $\mathrm{Pr}_{2}[\hat{\mathrm{~h}}=\mathrm{k}]$ functions obtained from the experimental dala of dichotic condition 2. The
$\sigma(f) / f$ functions have the typical U-shape which was also found in an earlier study [11], and have also the same general magnitude. The low-frequency slopes of these functions, however, are much steeper than the average slope found in that earlier study. We think this is caused by an over-estimate of $\sigma(f)$ at low frequencies in the present experiments. Partial frequencies below 1000 Hz , with fundamentals limited between 200 and 300 Hz , could occur only for very low harmonic numbers where identification
close to perfect and occasional mistakes are more made through carelessness or poor attention than through insufficient salience of pitches
(10)

Fig. 2. Model variance or "noise" functions" $\sigma(f) / f$ computed from the averaged functions $\mathrm{Pr}_{1}$ and $\mathrm{Pr}_{2}$ shown in Fig. 1b.


Fig. 3. Correct identification scores for both simultaneous fundamentals as a function of d defined by Eq. (5). (a) Diotic condition
1 (circles) and dichotic condition 2 (crosses). (b) Dichotic condi1 (circles) and dichotic condition 2 (crosses). (b) Dichotic condi-
tions 3 (circles) and 4 (crosses).

In order to examine the influence of spectral interference on pitch dentification performance, all 810 different stimuli were mapped on a frequency-difference measure $d$, defined as:

$$
\begin{equation*}
d=\sqrt{d_{1}^{2}+d_{2}^{2}}, \tag{5}
\end{equation*}
$$

where $d_{1}$ represents the smallest frequency difference between an two harmonics and $d_{2}$ the next smallest difference in the total of $5 \%$ and, in order to limit the general degrading effect of high harmonic numbers on pitch identification performance, only stim lif were included having lower harmonic numbers $m$ and $n$ of or less. Percentages correct pitch identification of both fundamentals as a function of $d$ are shown in Fig. 3 a for experimenta conditions 1 and 2 . Under diotic condition 1 values of $d$ below $(m+1) f_{01},(n+1) f_{02}$ must have interfered with one anothe because of limited frequency resolution in the cochlea. Under dichotic condition 2 such interference was not possible because potentially interfering partials went to different ears. Figure 3 a shows that only for the lowest $d$-values, between 0 and $5 \%$, ther is a noticeable difference between the scores of conditions 1 and condition 1 although degraded is still well above the expected chance level of $10 \%$ correct. Similar results were obtained with the data from conditions 3 and 4. They are shown in Fig. 3b.
Iv. MODEL SIMULATIONS

The data presented in the previous section show a general performance deterioration with increasing (lower) harmonic numbers $m$ and $n$, and also a dependence of performance on the the presen bation conditions 1 through 4. The data still provide insufficient information, however, about the relative contribution of parsing orspared with errors caused by interference of partials of mutual dependence of pitch identification processes. To study th infuence of parsing errors in more detail, a computer simulation experiment was performed with the model discussed in Sect. rived from the data of condition 2 were substituted in the model to specify the exact amount of noise to be added to each frequency component of the simulation input. For all 810 stimuli 25 compu tations were made (with new noise samples added to partials each time) of the maximum log-likelihood function of Eq. (2) withou Simulations were correct parsing, as ous in Case B of Sect. uli for which the correct parsing was always obtained were put in a stimulus subset PNS (parsing-non-sensitive). The remaining stimuli, for which (occasionally) the likelihood function cam out maximum with the wrong parsing, were put in the subset P (parsing-sensitive). With the subsets PS, PNS and also with the
entire set PS for all four stims, the simulation experiment was now repeated tion of the appropriate $\sigma(f)$ / f function obtained from the data of each particular subject under that condition with stimulus sub${ }^{\text {set PNS. Simulation was done with the use of Eq. (3), implying }}$ of ald 2 ge of the correct stimulus parsing, and with substitutio implying the permutations outlined under Case B of Sect. as a percentage correct this knowledge. The results, expressed and $f_{02}$ pooled over all values of $m$ and $n$ are shown in Fig 4 for stimulus conditions 1 and 2 . For each of the three stim



Fig. 4. Measured (solid lines) and simultated performance levis with (dashed lines) and without (dotted lines) knowledge of stimulus parsing. Columns PS are for the stimulus subset hat is induce such errors, the central columns for the entire set. Left: diotic condition 1 ; right: dichotic condition 2
ulus (sub)sets, the solid line represents the actual performance of the subject, the dashed line the performance level simulated with parsing knowledge, and the dotted line the level simulated without this knowledge. For the PNS-subset one expects all three efformance levels to be identical. The fact that this is not ex
 experimental and sumbers, but represents a small uncertainty hout the details of the simultated decision strategy. One also observes that for the subset PS and for the entire stimulus set PS+PNS the performance level of all subjects (solid lines) is much closer to the performance level simulated with parsing knowledge (dashed lines) than to the level simulated without hat knowledge (dotted lines). This is true for dichotic condition 2 as well as for ere obtained for stimulus conditions 3 and 4. This finding is mportant because in the diotic condition 1 no explicit parsing information was supplied to the subjects, and in conditions 3 and 4 an explicit attempt was actually made to supply them with false arsing information. If this wrong information had been use level, the subjects, their performance would have been at chance level, which above chance level for those conditions, however, as is evident from Figs. 1c,d. The empirical and simulated results tell is that subjects somehow do have a fairly accurate knowledge of the proper interpretation of the various stimulus partials in the percept of simultaneous complex tones, but that this knowledge is not obtained on the basis of ear inpul. an the basis oxplency to group perceived partials holistically on the basis of common fundamental. Something similar was also ound by Deutsch [14] and Butler [15] who used entirely differen musical paradigms.

## V. CONCLUSIONS

From the experimental and simulated results of this study the following conclusions are drawn:

1. The task of identifying two pitches when exposed to two simultaneous complex tones is separable into two pitch identification processes which are to a large extent independent. The small amount of mutual dependence that is sometimes observed tends to support the notion that the more salient pitch, represented by lower-order harmonics, is processed first and interferes with the processing of the less salient pitch. This mutual dependence of the two processes, small as it may seem, is largely responsible for the degradation of performance when going from dichotic condition 2 to diotic condition 1 and finally to dichotic conditions 3 and 4.
2. Information of the correct parsing and interpretation of perceived stimulus partials is, to a large extent, available to the central pitch processor. It is independent of the manner in which partials are distributed between the ears. This is consistent with other results on simultaneous-melody perception in the literature [14,15], and with informal observation of ordinary musical practice in which both ears are always exposed to all partials of simultaneously-playing musical instruments.
3. Interference of spectrally close partials has a surprisingly small effect on pitch identification for two complex tones, at least. as long as either tone is represented by harmonics of sufficiently low order. Although it is known that high-order harmonics of a single complex tone do not contribute much to fundamental pitch sensation [9] and are as such not available to the central processor [11], it now appears that aurally non-resolved harmonics belonging to different tone complexes are not entirely discarded. They may instead be transformed into a single (noisy) percept that can be used more than once by the processor when filling in the variables of Eqs. (1) or (2). This idea will be investigated further in a future study.
4. The human central pitch processor appears not to be hardwired or specifically programmed for one particular way of processing stimulus tones. On the contrary, its processing algorithm appears to be quite cooperative and interactive with the task it has to execute.

## VI. ACKNOWLEDGMENTS

This work, initiated under a grant from the U.S.A. National Institutes of Health (NS 11680) was executed with a grant from the Netherlands Foundation for the Advancement of Pure Research (ZWO 560-260-009).

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