A Probabilistic Model of Lexical and Syntactic Access and Disambiguation

DANIEL JURAFSKY

International Computer Science Institute, Berkeley

and

University of Colorado, Boulder

The problems of access—retrieving linguistic structure from some mental grammar—and disambiguation—choosing among these structures to correctly parse ambiguous linguistic input—are fundamental to language understanding. The literature abounds with psychological results on lexical access, the access of idioms, syntactic rule access, parsing preferences, syntactic disambiguation, and the processing of garden-path sentences. Unfortunately, it has been difficult to combine models which account for these results to build a general, uniform model of access and disambiguation at the lexical, idiomatic, and syntactic levels. For example, psycholinguistic theories of lexical access and idiom access and parsing theories of syntactic rule access have almost no commonality in methodology or coverage of psycholinguistic data. This article presents a single probabilistic algorithm which models both the access and disambiguation of linguistic knowledge. The algorithm is based on a parallel parser which ranks constructions for access, and interpretations for disambiguation, by their conditional probability. Low-ranked constructions and interpretations are pruned through beam-search; this pruning accounts, among other things, for the garden-path effect. I show that this motivated probabilistic treatment accounts for a wide variety of psycholinguistic results, arguing for a more uniform representation of linguistic knowledge and for the use of probabilistically-enriched grammars and interpreters as models of human knowledge of and processing of language.

1 INTRODUCTION

The problems of access—retrieving linguistic structure from some mental grammar—and disambiguation—choosing among combinations of these structures to correctly parse ambiguous linguistic input—are fundamental to language understanding. Recently, a number of computational cognitive models of these tasks have appeared, including models of lexical access,

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Correspondence and requests for reprints should be sent to Daniel Jurafsky, Department of Linguistics, University of Colorado at Boulder, Boulder, CO 80302-0295.
(Cottrell, 1985; McClelland & Elman, 1986), lexical disambiguation (Small & Rieger, 1982), syntactic disambiguation (Abney, 1989; McRoy & Hirst, 1990; Shieber, 1983), the access and disambiguation of idioms (van der Linden, 1992; van der Linden & Kraaij, 1990; Wilensky & Arens, 1980), and the difficulties in disambiguating garden-path sentences (Gibson, 1991; J. Henderson, 1994; Pritchett, 1988).

It has proven difficult to combine these models to build a general, uniform model of access and disambiguation at the lexical, idiomatic, and syntactic levels. For example, psycholinguistic theories of lexical access and idiom access and parsing theories of syntactic rule access have almost no commonality in methodology or coverage of psycholinguistic data. In part, this has been due to an assumption that these levels, or at least the lexical and syntactic levels, are functionally distinct in the human language processor. In part, a uniform model has been difficult to reconcile with traditional linguistic models, which have tended to split into phonological, syntactic, and semantic modules. And finally, uniformity has met with problems from the psycholinguistic data which have suggested quite different models for lexical and syntactic processing. For example, there is a great deal of evidence for parallelism in lexical processing (Swinney, 1979; Tanenhaus, Leiman, & Seidenberg, 1979; Tyler & Marslen-Wilson, 1982). Thus, in most models of lexical access, multiple candidate words are activated from the mental lexicon based on orthographic or phonological input. Psycholinguistic models of syntactic processing, however, have generally been serial, rather than parallel. That is, only one parse of the input is maintained at all times. One reason for this has been the need to account for the unacceptability of garden-path sentences like (1).

(1) # The horse raced past the barn fell. (Bever, 1970)

As we will see in §4, sentences like (1) are difficult because they are locally ambiguous at some point in the processing (e.g., between a main-verb and reduced-relative reading of the word 'raced'). In most syntactic accounts of this data, the serial nature of the parser requires only one of these interpretations to be chosen, and local heuristics cause the incorrect main verb reading to be chosen. These serial models of syntactic processing are incompatible with the parallel models of lexical processing.

Finally, while robust frequency effects have long been noted in lexical processing, neither linguistic theories of syntax, nor the parsing algorithms which use them to build parses, have suggested methods for dealing with probabilistic effects in syntax.

Despite these problems and differences, most researchers agree that at least some features of the language processing architecture are uniform across levels. For example there is evidence for on-line processing at the lexical, idiomatic, and syntactic levels, including evidence from comprehension

But more fundamentally, recent years have seen a convergence among linguists and psychologists toward less modular, more interactive models of language and language processing. The rise of unification-based linguistic theories (Bresnan, 1982; Pollard & Sag, 1987; Sag et al., 1985) has led to the ability to represent information at every level of representation with the same formalism. Tanenhaus and associates (Spivey-Knowlton, Trueswell, & Tanenhaus, 1983; Trueswell & Tanenhaus, 1991; Tanenhaus et al., 1989) have argued that semantic aspects of lexical information such as verbal semantic valence play an early role in syntactic parsing, and that semantic information plays a role in garden-path effects. Finally, MacDonald (1993) and Kawamoto (1993) have noted a number of similarities between lexical and syntactic disambiguation, and suggested that a uniform constraint-based mechanism might be built to account for disambiguation across linguistic levels.

In this article we follow these recent directions to propose a single computational architecture that models both access and disambiguation, accounting for psycholinguistic results on lexical and idiomatic access, lexical and syntactic disambiguation, and garden-path sentences. The model is based on a preference for coherence in interpretation. Intuitively, constructions are more likely to be accessed when they are more coherent with the previous context, and interpretations are preferred over other interpretations when they are more coherent. Significantly, we show that this intuition of coherence can be formalized with a probabilistic foundation. While linguistic grammars which include frequency information date at least back to Ulvestad’s (1960) work on subcategorization, only recently has a body of techniques become common for dealing with stochastic grammars (Baker, 1979; Fujisaki, Jelinek, Cocke, & Black, 1991; Jelinek & Lafferty, 1991) in which every rule of the grammar is augmented with a conditional probability. Recent studies have applied these probabilities to linguistic theories like LTAG (Resnik, 1992), and have shown the use of such probabilistic grammars for modeling synchronic change (Tabor, in press) and learning (Stolcke, 1994; Stolcke & Omohundro, 1993). In this article we show that a probabilistic model differs from the frequency-based models traditional in psycholinguistics and argue that true probabilities are essential for a cognitive model of sentence processing.
The algorithm is designed in the framework of Construction Grammar (Fillmore, Kay, & O'Connor, 1988; Goldberg, 1991; Goldberg, 1992; Kay, 1990; Koenig, 1993; Lakoff, 1987; Lakoff, 1993; Lambrecht, in press), a unification-based theory in which the mental lexicon, idiom list, and grammar are represented as a uniform collection of grammatical constructions. Construction grammar is a sign-based theory, like HPSG (Pollard & Sag, 1987), in which each construction represents well-formedness conditions across various domains of linguistic knowledge, including phonological, morphological, syntactic, semantic, and pragmatic domains. Although we describe the algorithm in the construction grammar framework, it can be applied to any sign-based unification-based linguistic theory which allows the expression of predicate valence (HPSG, cognitive grammar, Montague Grammar, categorial grammar, RRG, probably LFG) simply by the addition of the appropriate probabilities.

The model consists of a parallel parser augmented with probabilities and a pruning heuristic which models both access and disambiguation. The model addresses access by unifying lexical access, syntactic rule access, and idiom access. In this Bayesian access or evidential access algorithm, phonological, syntactic and semantic evidence, top–down as well as bottom–up, is integrated to determine which constructions to access, by computing the conditional probability of the construction given each piece of evidence. Lexical, idiomatic, or syntactic constructions are activated in parallel when the weight of evidence passes a threshold. Our model of construction disambiguation, the local coherence model, similarly conflates lexical disambiguation, idiom disambiguation, syntactic disambiguation, and semantic disambiguation. The algorithm again computes the conditional probability of each interpretation given the input, and ranks interpretations according to these probabilities. The algorithm prunes these interpretations to keep the search space manageable; a by-product of this pruning is the garden-path effect, which we model as the pruning of what turns out to be the correct parse.

The actual probability computation involves two sources of expectations, constituent and valence information. Constituent expectations arise from the phrase-structural skeleton of the grammar. Each context-free rule is annotated with a conditional probability that the left side will expand to the right side. These probabilities place a distribution over the kind of daughters each phrasal node requires. Valence, or subcategorization expectations are associated with verbs and other predicates. Each predicate may express constraints on the syntactic, thematic, or semantic form of their arguments. For each such argument, the predicate is also annotated with a conditional probability that the argument will be filled.

We draw three conclusions from this work. First, we show a unified model of a number of psycholinguistic phenomena, including parse preferences, explaining the difficulty of garden-path sentences, and results on lexical and
idiom access. The model requires only simple probabilistic augmentations
to concepts of phrase structure and valence that already exist in most
modern grammars and theories of sentence processing. Second, we argue
that these results demonstrate the need for probabilistic augmentation to
grammatical theories if they are to be able to deal with a broad range of
phenomena as a processing model. In this vein our model is part of a larger
effort involving computational implementations of linguistic theory (Jurafsky,
1992) and large-scale implementations of probabilistic augmentations to lin-
guistic models for speech recognition (Jurafsky et al., 1995; Tajchman et
al., 1995).

Finally, the idea that lexical, idiomatic, syntactic, and semantic struc-
tures are uniformly represented and processed contrasts sharply with the
traditional modular view, which holds that each of these kinds of linguistic
knowledge is separately represented and processed. A number of linguists
and psychologists have recently argued that an integrated view of linguistic
structure and processing is necessary to account for linguistic data (Fillmore
et al., 1988; Goldberg, 1991; Lakoff, 1987; Lambrecht, 1995; Langacker,
1987; MacDonald, 1993; McClelland, St. John, & Taraban, 1989; Spivey-
Knowlton & Sedivy, in press; Spivey-Knowlton et al., 1993; Taraban &
McClelland, 1988; Trueswell & Tanenhaus, 1991). We hope that by showing
that a single parallel probabilistic algorithm deals uniformly with lexical
access data, idiom processing, parsing preferences, and garden-path data,
to provide additional evidence that a uniform, nonmodular theory of lan-
guage representation and process is possible.

2 ARCHITECTURE

Access and disambiguation are only a small part of a complete model of
parsing. In order to maintain the advantages of computational efficiency in
the face of complex recursive structure, we assume our parser is built with
the standard dynamic programming (chart) parsing architecture. Then our
access and disambiguation algorithms are implemented as a set of pruning
heuristics that augment this standard algorithm. In the next sections we
give an overview of the motivation for using pruning heuristics as a cogni-
tive model; detailed arguments will be presented in §3 and §4. We then turn
to our model of linguistic representation.

2.1 Parallel Parsing

As we will discuss in §3, psycholinguistic results on lexical access and idiom
access, as well as recent results on syntactic processing and our own argu-
ments in §4 on garden-path processing lead us to believe that the underlying
architecture of the human language interpretation mechanism is parallel
(although cf. Frazier [1987]).
However, many researchers have noted that without some special attempts at efficiency, the problem of computing parallel syntactic structures for a sentence can be quite complex. Church and Patil (1982) showed, for example, that the number of ambiguous phrase-structure trees for a sentence with multiple preposition-phrases was proportional to the Catalan numbers, while Barton, Berwick, and Ristad (1987) showed that the need for keeping long-distance agreement information and the need to represent lexical ambiguity together make the parsing problem for a grammar that represents such information NP-complete. If we expect our parsing algorithms to extend to spoken-language input, the ambiguity problem explodes even further, first, because just allowing the input to be strings of phones rather than words adds phonetic parse and lexical segmentation ambiguity, and second, because phonetic estimation itself is likely to require probabilistic, nondiscrete inputs to the parsing algorithm.

All solutions to the problem of maintaining multiple parses of an ambiguous sentence, while still parsing in polynomial time, involve dynamic programming techniques. These methods, from the well-formed substring table (WFST) of Kuno (1965), to the chart parsing algorithm of M. Kay (1973) and the Earley algorithm of Earley (1970), all essentially trade memory for processing time. The parser stores the common subparts of multiple parses, allowing subparses to be represented only once, instead of once per parse tree.

We assume that the human parser will need to use some such dynamic programming algorithm for parsing control and for rule-integration.

2.2 Why Pruning?

However, the efficiency gained by dynamic programming algorithms for pure syntax may not generalize to the problem of interpretation. For example, if two parse trees both include an NP, the dynamic programming algorithm can simply store the NP once, because the internal structure of the NP is irrelevant to the global parse. But if two interpretations share the same NP, it may not be possible to store the NP only once, because its internal structure, and particularly its semantic structure, is relevant to the interpretation, and may be needed by the interpreter to produce part of an on-line interpretation. Building the semantics of the NP into the interpretation may involve binding variables differently in the context of different interpretations. Although some semantic structure can most likely be shared, the sharing will not be as efficient as for syntactic structure.

However, there is another simple way to improve the computational tractability of parsing. The results of Church and Patil (1982) and Barton et al. (1987) rely on the fact that syntactic ambiguities in these parsers are not resolved until after the entire sentence has been parsed. It is the need to represent ambiguities for indefinite lengths of time in parsing that causes complexity. One way for a cognitive model of parsing to be computationally
tractable, then, is to do some sort of on-line pruning. That is, rather than searching the entire space of possible interpretations for an input utterance, we rank our hypotheses and abandon the disfavored ones, or indeed find some way to avoid even suggesting (accessing) them.

Our models of access and disambiguation are based on this idea. As we will see in §3, in a standard model of parsing like bottom-up parsing, every possible rule which could possibly account for the input data is accessed. As the rule-base grows, this becomes quite a large set. If we would like our access algorithms to extend to speech processing, the nondeterminism of phonetic input makes the number of possible structures which could be input even larger. In order to make the access problem tractable, we argue that it will need to do some sort of pruning, either by accessing many structures and then dropping some, or by accessing fewer structures in the first place.

While access pruning happens as linguistic structures are suggested from the mental grammar, disambiguation happens after structures are integrated together. Disambiguation is a kind of pruning by definition, selecting an interpretation which is better than other interpretations on some metric. In order to avoid the computational complexity of maintaining ambiguities over long distances, our model disambiguates on-line by pruning unlikely interpretations.¹

Both the access and disambiguation pruning algorithms work by using probability to rank constructions to be accessed or interpretations to be disambiguated. In the next section we introduce the grammatical formalism that will allow these probabilities to be computed.

2.3 Grammar: Uniformity and Probability
The representational component of our processing model is based on construction grammar. Construction grammar adopts from traditional grammar the idea that a grammar consists of a large collection of grammatical constructions. Each grammatical construction is a sign, a conventionalized association between form and meaning. Formally, these signs are represented as typed unification-based augmented context-free rules in the version of construction grammar we consider here. In these senses construction grammar resembles most phrase-structure and unification-based linguistic theories, including among others LFG (Bresnan, 1982), HPSG (Pollard & Sag, 1987),²

¹ See Mathis and Mozer (1995) for other arguments for the computational utility of disambiguation.
² Note that uniform processing of lexical and syntactic knowledge can take place whether we assume that lexical entries or CF rules are the more fundamental mechanism; i.e., lexical entries can be modeled as trivial lexical insertion rules in a CFG formalism, or alternatively phrase structure constituency can be represented with special attributes with complex feature structures.
and certain versions of Categorial Grammar. Other versions of construction
grammar, and closely related theories like Cognitive Grammar (Langacker,
1987), share the sign-based foundation of the version of construction gram-
mmer we discuss here, but differ on the use of feature structures and CF rules
as a formal implementation. Like all these similar approaches, construction
grammar contrasts with the nonsign-based Principles and Parameters/
Minimalist approach to grammatical representation. This article will also
show how to augment construction grammar with certain kinds of probabil-
ities. As we will see, these same probabilistic augmentations could be made
to any other of these sign-based theories which meets certain requirements.

Our processing model relies on four assumptions concerning grammatical
representation:

1. The representation of constituent structure rules as mental objects.
2. A uniform context-free model of lexical, morphological and syntactic
   rules.
3. Valence expectations on lexical heads.
4. A lack of empty categories.

In the rest of this section we describe and motivate each of these assump-
tions, discussing the extent to which each is true of construction grammar
and other theories.

The first assumption deals with what Fillmore et al. (1988) refer to as
"the distinction between knowing and figuring out," what is often called in
psychology the 'economy of representation' hypothesis. L. Henderson
(1989) has remarked that linguistics and psychology hold quite different
views about representational economy. A fundamental motivating principle
for many linguistic theories is the minimal-redundancy, maximally eco-
nomical grammar or lexicon. Psychological models, on the other hand, often
emphasize the vast storage capability of the mind. Construction gram-
mar borrows from these tendencies in psychology by modeling grammar as
a comparatively large collection of structures, including larger phrase-
structure rules as well as lexical entries. Generalizations across these struc-
tures, including simple lexical entries as well as complex phrase-structural
projections such as the correlative conditional construction of Fillmore
(1986) described as follows, are captured by a type hierarchy, resembling
the type hierarchy of HPSPG or the schematic networks of cognitive gram-
mar. Figure 1 shows part of the construction grammar type hierarchy.

The idea that the grammar is a collection of these structures arranged in a
type hierarchy has two implications for processing. First, because construc-
tions and the type hierarchy are the sole mechanism for capturing generali-
zations, the theory places no requirement on the processing mechanism to
implement redundancy rules, metarules, movement, or enrichment of surface
form. We will use the lack of traces to model the psycholinguistic results of Tanenhaus, Stowe, and Carlson (1985) in §4. In as much as LFG has eliminated traces, and recent versions of HPSG have begun to do away with traces (Pollard & Sag, 1994) and lexical rules (Krieger & Nerbonne, 1993), this assumption of representational minimalism is met by a number of theories. Second, allowing these structures means the theory has an independent, relatively rich phrase-structural component, in which phrase-structural patterns are stored in the mental inventory. We will see the need for storing constituent-structure patterns (and their probabilities) in modeling garden-path effects in §4. Besides construction grammar, the storage of these structures is assumed by the sign-based theories (HPSG, Cognitive Grammar) and LFG, but not by categorial grammar, and also not in the PP/Minimalism paradigm, in which phrase structural categories act only as a kind of principle-based filter on possible trees, and have no independent existence in any kind of mental store (and hence cannot be augmented with probabilities).

The second fundamental assumption of construction grammar is representational uniformity. Uniformity has a number of facets. First, construction grammar makes no distinction between the lexicon, the morphology and the syntactic rule base; each is represented with an augmented CFG formalism. A lexical entry is a construction that may have only a single constituent. Morphological knowledge is represented with context-free rules (Selkirk, 1982) for example, an affix like English plural 's' may be treated as a rule N[ + plural] → STEM 's'\(^4\). This assumption allows our processing model to treat lexical and syntactic processing equivalently; although we don’t consider morphology in this paper, we expect to apply the same parsing algorithms

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\(^3\) For arguments that inheritance together with on-line type construction is sufficient to replace lexical rules, see Koenig and Jurafsky (1995).

for morphology as we have for syntax. The idea that lexical, morphological, and syntactic structures are uniformly represented is a fundamental principle of sign-based linguistic theories such as Cognitive Grammar and categorial grammar. With the exception of morphological knowledge this is also true of HPSG. Because in sign-based theories like construction grammar constructions can be associated with semantics like lexical items, the access and disambiguation processes can make reference to semantic structure.

As an example of a construction and the use of inheritance to capture generalizations, consider the CORRELATIVE CONDITIONAL construction described by Fillmore (1986) and (1988) and McCawley (1989), which models sentences like (2):

(2) a. HERMIA: The more I hate, the more he follows me.
    b. HELENA: The more I love, the more he hateth me.

Figure 2 shows Fillmore's analysis of the construction.

Note that the correlative conditional construction is a subtype of the conditional construction, which is itself a subtype of the subordination construction. In addition, the second and fifth constituents are constrained to be of type COMPARATIVE. Fillmore argues that sentences like (2) simultaneously exhibit conditional and comparative properties, and that these various grammatical properties of the construction are inherited from the conditional and comparative constructions.

Although LFG makes a crucial distinction between the lexicon and the grammar, recent LFG work has discussed similarities between lexical and syntactic structure such as assigning semantic rules to specific syntactic constructions like relative clauses (Dalrymple, 1992).
For example, the correlative conditional inherits from the conditional construction many facts about the tense and aspect possibilities of the two clauses, such as the suppression of future will in the protasis:

(3)  a. The faster you (*will) drive, the sooner you'll get there.
     b. If you (*will) drive fast, you'll get there by 2:00.

As McCawley (1989) shows in (4), the second and fifth constituents of the correlative conditional exhibit the complete range of ordinary comparative morphology; -er on short stems, more on long stems, and suppletion with good, bad, will, badly, much, many, little.

(4)  a. The worse the weather gets, the happier I am that we stayed home.
     b. The worse he behaved, the less attention we paid to him.
     c. The better you treat him, the less trouble he'll give you.

The correlative conditional has a number of idiosyncratic properties which cannot be predicted from its supertypes. For example, some particulars about the use of the comparative are unique to this construction; the two comparatives ("the more time" and "the less") are fronted and semantically express a correlation between an independent and a dependent variable. For a sentence like (5), the interpretation might be paraphrased as "The degree to which I spend time on this problem will determine the degree to which I understand it".

(5)  The more time I spend on this problem, the less I understand it

Again, a guiding intuition of construction grammar is that such complex, nonlocal constructions are stored along with lexical entries in the mental grammar.

We turn now to the probabilistic augmentations we propose for constructions. We argue that constructions should be augmented with probabilities in two ways. First, every construction will be annotated with a simple probability. These express the prior probability of choosing a particular construction at random from the mental lexicon or from a random utterance. This prior probability resembles the "resting activation" of the frequency-based models traditional in psychological theories of lexical access. Second, constructions are augmented with probabilities in every case where some part of the construction expresses a linguistic expectation. A feature-based lexical grammar expresses expectations in at least two ways. First, each context-free rule expresses an expectation for the constituents which constitute the rule. That is, given that we have evidence that the parser has seen a certain phrase structure rule, we can expect each constituent of the rule to appear. Second, each valence-bearing predicate expresses an expectation for its valence-fillers. That is, given evidence for a predicate, we can expect to find further in the string a constituent which will satisfy each valence (or
subcategorization) slot of the predicate. Our model allows each of these expectations to be probabilistic.

How are these probabilities to be defined? The prior probability of a construction is the simplest. We assume here a frequentist interpretation of prior probabilities. That is, the prior probability of a construction is just a maximum likelihood estimate from relative frequencies. We make the simplifying assumption that these frequencies can be computed either from psychological norming studies or from carefully chosen corpora. The most widely available balanced corpus is the million word Brown corpus. Francis and Kučera (1982) give frequencies for every lexical item in the Brown corpus, and also for simple lexical categories. Computing frequencies for phrasal constructions requires parsing the corpus. For phrasal construction frequencies we used the Penn Treebank (Marcus, Santorini, & Marcinkiewicz, 1993), which includes a parsed segment of the Brown corpus.

Given the prior probabilities for each construction, computing the probabilities for constituent expectations can be done with a simple normalization step. Constituent expectation probabilities are in fact the standard Stochastic Context-Free Grammar probability. This is the conditional probability of the right hand side of a rule, given the left hand side, that is, the probability of a particular expansion of a left-hand side. Thus, if the expansions of $A$ consist of SCFG rules of the form

(6) $[p_i] A \rightarrow B C$

(7) $[p_i] A \rightarrow D E F$

then $p_i$ is the probability that $A$ will expand to $BC$, while $p_i$ is the probability that $A$ will expand to $DEF$. Since these probabilities are all conditional on $A$, it follows that $\Sigma_i p_i = 1.0$. Thus, we can compute the SCFG probability of any rule just by dividing the prior probability of the rule by the sum of the prior probabilities of all expansions of the left-hand side.

For example, in order to compute the probability of the simplified rule (8), we would use the treebank to get a frequency for all NPs (52,627), and then for those NP's which consist of a Det and an N (33,108). The conditional probability is then $33,108/52,627 = .63$.

(8) [.63] $NP \rightarrow Det N$

The second kind of grammatical expectations are valence expectations that a lexical predicate bears for its arguments. As discussed above, the valence of a predicate in construction grammar may constrain the syntactic category, grammatical function, thematic role, or semantics of their arguments. Each lexical predicate may bear a valence attribute which specifies a set of valence arguments. In our probabilistic model, each of these arguments is augmented with a probability of occurrence. This probability expresses the conditional probability of the argument being filled given the
predicate. Obligatory arguments, for example, will have unity valence probability, while optional arguments will have a value between 0 and 1.

An alternative algorithm might assign probabilities to entire thematic grids rather than individual arguments. Although we express our probability computations in terms of individual-argument rather than thematic-grid probabilities, we will show our pruning algorithm is compatible with either model. The data we discuss in this article does not allow us to definitively choose between the two models of valence probability.

We have relied on two methods of determining valence probabilities. Connine, Ferreira, Jones, Clifton, and Frazier (1984) provide frequencies for the different syntactic subcategorizations of 127 verbs. These were determined by giving subjects a list of verbs and asking them to write a sentence for each one. The authors then computed frequencies on the corpus this produced.

For verbs which were not studied by Connine et al. (1984), we used the Penn Treebank (Marcus et al., 1993) just as for construction probabilities. That is, to determine the valence probabilities for a verb, we counted the number of times the verb occurred with each frame, and normalized by the total count for the verb. Since these initial sources of valence probabilities only include syntactic information, we have left the use of semantic valence probabilities for future research (see §5).

3 ACCESS

Every model of parsing, lexical access, or sentence processing includes a theory of access, and yet the problem of access has been strikingly balkanized. In this section, we compare previous models of access at the lexical, idiomatic, and syntactic levels, and hold them up to psycholinguistic results. We show that each of the models fails to model all the relevant data, and then present the Bayesian access model, based on computing the conditional probability of a construction given top-down and bottom-up evidence. We show that the model is consistent with psycholinguistic results and outline the model’s predictions.

3.1 Previous Models of Access

Consider the difference between serial and parallel access algorithms. The earliest models of lexical, idiomatic, and syntactic access were all serial, including lexical models like ordered-access (Hogaboam & Perfetti, 1975), models of idiom access like the idiom list hypothesis (Bobrow & Bell, 1973) and the direct access hypothesis (Gibbs, 1984), and models of syntactic access in parsers like PARSIFAL (Marcus, 1980) and the Sausage Machine (Frazier & Fodor, 1978). In serial models a single lexical entry is accessed from the lexicon at a time. For example, in the idiom list hypothesis, idioms are stored
in a separate idiom dictionary, and a single idiom is accessed when the computation of literal meanings for a string fails, while the direct access model of idiom access proposes just the opposite: idiomatic meaning is accessed first, and literal meaning is only computed when the access of idiomatic meaning fails.

In recent years, serial models of lexical and idiom access have fallen somewhat out of favor, due to extensive results which argue for parallelism (Swinney, 1979; Swinney & Cutler, 1979, etc). Thus, most researchers on idiom processing assume some form of the lexicalization hypothesis (Swinney & Cutler, 1979), which proposes that idioms are stored as long lexical items and accessed in parallel with normal lexical items.

Syntactic access is treated very differently in the psycholinguistic and computational communities. Psycholinguistic parsers (such as the early models of Bever, 1970, and Kimball, 1973, as well as later models like the Sausage Machine), are traditionally serial, motivated as they are by the garden-path results we will discuss in §4. Computational parsing models (such as the Earley, CKY, and chart parsers) are usually based on some form of dynamic programming algorithm, and hence are fundamentally parallel in nature.

However, we will argue in §4 following Kurtzman (1985), Norvig (1988), and Gibson (1991), that garden-path phenomena can be modeled with a parallel architecture as well, eliminating one of the strongest arguments for serial parsing. Although psycholinguistic results on syntactic access seem to be somewhat contended, a number of recent results argue for parallelism in syntactic processing as well (Boland, 1991; Gorrell, 1987; Gorrell, 1989; Kurtzman, 1985; MacDonald, 1993). Because all modern models of lexical access, and most models of idiom or syntactic access are parallel, then, the remainder of this section focuses on parallel access algorithms, examining previous models from two perspectives: how to access (i.e., what sorts of evidence to use), and when to access (i.e., the time course of access). For example, consider the problem of accessing either the syntactic rule \( S \rightarrow NP \ VP \) or the lexical item 'about' \( (\lambda \ b \ a \ w \ t) \) given some input. Figure 3 sketches four standard rule access models, showing the type of evidence each algorithm would use to access the rule \( S \rightarrow NP \ VP \) or 'about' \( \rightarrow \lambda \ b \ a \ w \ t \).

In the bottom-up model, access of a rule depends only on evidence from the input, and is delayed until the entire set of constituents has been seen. One interpretation of the lexicalization hypothesis (Swinney & Cutler, 1979) would process idioms in this way; van der Linden and Kraaij (1990) built a computational model of this proposal, in which idioms are not accessed until the entire idiom has been seen. In a top-down model (such as the selective access model of lexical access (Glucksberg, Kreuz, & Rho, 1986; Schvaneveldt, Meyer, & Becker, 1976), access depends only on previous expectations. Neither the top-down nor bottom-up models meets our concern with psycho-
PROBABILISTIC MODEL OF ACCESS AND DISAMBIGUATION

Logical plausibility. For example, we cited in the introduction a large number of studies showing that language processing is strictly on-line, ruling out a bottom-up model which delays until every constituent has been seen. Similarly, a number of studies have shown results inconsistent with selective access and other purely top-down models (Swinney, 1979).

The left-corner model combines advantages of the bottom-up and top-down models; a rule is accessed by only the first constituent, and then processing continues in a top-down manner from the rest of the rule. Such models have been proposed for both syntactic parsing and lexical access. For example, Marslen-Wilson’s (1987) Cohort model of lexical access in speech perception is in many respects a left-corner model, using bottom-up information to access entries, and then top-down information to process them further. In the Cohort model, bottom-up phonetic information is used to access a set of lexical entries whose initial phonemes are compatible with the input so far. The set, called the cohort set, is then weeded out by using various top-down as well as bottom-up information sources to prune words which don’t meet their constraints.

The final class of models, the island models, propose even more sophisticated ways of accessing a rule. In head-corner access, only the head of the first constituent need be seen to access a rule. While the head-corner model was proposed independently by van Noord (1991) and Gibson (1991) for syntactic parsing, Cutler and Norris’s (1988) MSS model of lexical segmentation is essentially a head-corner model applied to speech. The MSS model accesses words based on their first stressed syllable; the stressed syllable thus acts as the parsing head.

Finally, in what can be viewed as an extension to the head-corner model, two algorithms (Cacciari & Tabossi, 1988; Wilensky & Arens, 1980) have been proposed which mark specific constituents idioms as the key or indexing clue, and access idioms only after this constituent is seen. This allows these algorithms to model results indicating that the access of different idioms will occur at differing positions in the idiom.
3.2 Problems with Previous Models

Clearly there is a trend in more recent access models to be more and more sophisticated about the kind of evidence that is needed to access a rule or structure. Unfortunately, none of these models is quite sophisticated enough, and all suffer from two major problems. The first is their inability to handle timing effects; in particular, construction-specific, context-specific, and frequency effects in access. The second is their reliance on a single kind of information to access rules; either strictly bottom-up or top-down information, strictly syntactic information like the restriction algorithm of Shieber (1985), or solely semantic information, as in conceptual analyzers like Riesbeck and Schank (1978) or in Cardie and Lehnert (1991) or Lytinen (1991). Thus, each of these models is only able to model a particular range of evidential effects.

Consider the psycholinguistic evidence on timing. First, there is evidence that access timing is different for different constructions. The access point (point in time when the construction is accessed) for different constructions may be quite distinct. For lexical constructions, Tyler (1984) and Salasoo and Pisoni (1985) show that while the average access point for lexical items is approximately 150 ms after word-onset, timing is quite dependent on the frequency of the lexical item. High-frequency lexical items have higher initial activation than low-frequency ones (Marslen-Wilson, 1990), are accessed more easily (Tyler, 1984; Zwitserlood, 1989), and reach recognition threshold more quickly (Salasoo & Pisoni, 1985; Simpson & Burgess, 1985). Swinney and Cutler (1979) showed that some idioms were not accessed immediately after the first content word of the idiom, but rather that it took at least two words to access the idiom. Cacciari and Tabossi (1988) found that different idioms are accessed at quite different rates.

One way to model the access timing differences between constructions is with the island or key algorithm of Cacciari and Tabossi (1988) and Wilensky and Arens (1980). In these algorithms, each construction in the grammar would have one of its constituents (the key or island) marked. The construction could only be accessed after the key had been seen. One problem with this approach is that it requires specifying this information for each construction in the language; an approach which was able to predict the access point without having to learn it for each construction would be preferable.

A more serious problem with the island/key algorithm is its inability to account for the second class of timing results. These results show that different contexts change the access point even for the same construction, that is, that the access point is context-sensitive. Cacciari and Tabossi (1988) showed that the access of idioms was faster in the presence of context. Salasoo and Pisoni (1985) showed the same for lexical entries. Marslen-Wilson et al. (1988) showed the dual case—that anomalous contexts can slow down the access point of lexical constructions, and that the more anomalous the contexts, the higher the response latencies.
Thus, whatever sort of access algorithm we propose, it must allow the accumulation of evidence to cause some constructions, in some contexts, to be accessed faster than other constructions, in other contexts.

3.3 The Probabilistic Model

Our proposal is that access and disambiguation should be treated with a single probabilistic model. We refer to the access part of the model as the Bayesian or evidential access algorithm. For each construction, we compute the conditional probability of the construction given the evidence. Evidence can come from syntactic, semantic, and lexical sources, both top-down and bottom-up. Constructions are accessed according to a beam-search algorithm. In beam-search, every construction falling within a certain percentage of the most highly-ranked construction is accessed. We propose that this beam-width, which we call the access threshold $\alpha$, is a universal constant in the grammar.

For each construction, the conditional probability of a construction given top-down evidence is relatively simple to compute in a Construction Grammar or any other augmented-stochastic-context-free formalism. Recall that the SCFG prior probability gives the conditional probability of the right hand side of a rule given the left hand side.

Given these probabilities, the conditional probability of a construction $c$ appearing given some top-down evidence construction $e^+$ can be computed from the probability that $e^+$ will expand to $c$. In particular, since the parser operates left to right, the top-down probability $P(c|e)$ is the probability that the evidence construction $e^+$ left-expands to $c$:

$$P(e^+ \overset{L^*}{\rightarrow} c)$$

In a context-free grammar, a nonterminal $a$ left-expands to a nonterminal $b$ if there is some derivation tree whose root is $a$ and whose leftmost leaf is $b$. Consider the toy example in (10).

$$
\begin{align*}
(10) & \quad \text{a. } [.5] \quad S \rightarrow NP \ VP \\
& \quad \text{b. } [.7] \quad VP \rightarrow V \ NP \\
& \quad \text{c. } [.2] \quad VP \rightarrow V \ NP \ PP \\
& \quad \text{d. } [.1] \quad VP \rightarrow Adv \ VP \\
& \quad \text{e. } [.1] \quad V \rightarrow 'eat' \\
& \quad \text{f. } [.3] \quad V \rightarrow 'go' \\
\end{align*}
$$

Suppose the parser is in the process of parsing a sentence and has found the first NP of an S. The left expansion of the next constituent, the VP, includes the symbols V, eat, Adv, and go. Figure 4 shows a schematic example of left expansion.

Equation 9 holds whether the evidence is valence evidence, that is, evidence from a lexical head that a particular complement (syntactic or semantic) is
expected, or constituent evidence, that is, evidence from an incomplete construction. If no recursive production cycles occur in the expansion of $e$ to $c$, then $P(c|e)$ is very simple to compute as the product of the probabilities associated with each rule in every possible expansion. If recursive production cycles occur, Jelinek and Lafferty (1991) give an algorithm for efficiently computing left-corner probabilities, while Stolcke (1995) shows how the standard Earley parser can be augmented with them. Jurafsky et al. (1995) shows an application of a version of this algorithm as a model of lexical access for a speech recognizer.

In order to compute the conditional probability of a construction $c$ given evidence $e$ which is bottom-up evidence, we use the Bayes Rule:

$$P(c | e) = \frac{P(e | c)P(c)}{P(e)}$$

The probabilities $P(c)$ and $P(e)$ are simply the prior probabilities which annotate each construction, computed by normalizing frequencies. For the likelihood $P(e | c)$, we can now use the standard algorithm above in reverse, computing the probability that $c$ left-expands to $e$.

Once given these estimates for top-down and bottom-down probabilities, combining the evidence to estimate the posterior probability of a construction is quite complex. In current work, we are examining Bayes net realizations of the combination problem, which is in fact the subject of a large Bayes Net literature (Pearl, 1988). Certain assumptions may make the combination problem much simpler. For example, if we assume that top-down and bottom-up evidence interact disjunctively, they can be combined with a simple noisy OR-gate. Alternatively, we can make the simplifying assumption that $e^+$ can affect $e^-$ only through $c$, producing the following:

$$P(c|e) = P(c|e^+, e^-) = \frac{P(c, e^-|e^+)}{P(e^-|e^+)}$$

$$= \frac{P(c|e^+) P(e^-|c, e^+)}{P(e^-|e^+)}$$

$$= \frac{P(c|e^+) P(e^-|c)}{P(e^-|e^+)}$$
Since, as we show below, we will be comparing the ratios of the probabilities for different constructions, the denominator $P(e^-|e^+)$ remains constant and we can simply treat it as a normalizing factor, producing a very simple equation for evidence combination:

$$(12b) \quad P(c|e) = \lambda P(c|e^+) P(e^-|c)$$

However, psycholinguistic results on the relation between top-down and bottom-up evidence are quite controversial. While most studies agree that top-down evidence is used in parsing, many researchers have argued that bottom-up evidence is used first or more strongly (Frazier, 1989; Mitchell, 1989; although cf. MacDonald, 1993). This might be modeled by parameterizing the amount of weight given to top-down versus bottom-up evidence, perhaps via exponential weighting of the probabilities. A more complete investigation of evidence combination awaits future work.

Now given that the algorithm assigns a conditional probability to each construction, how should we choose which ones to access? We might choose to access any construction with nonzero probability. But as we argued in §2.2, accessing every possible construction may lead to far too many constructions, especially with noisy input such as speech. Furthermore, the results of Cacciari and Tabossi (1988) previously discussed argue that idioms are often not accessed until more than one word of them has been seen. This suggests that the access algorithm operates with some sort of pruning. For example, if the phonetic evidence suggests both a lexical entry and an idiom, but the lexical entry was much more probable, a pruning algorithm might keep the idiom from being accessed. Since the data rules out any sort of absolute threshold, our proposal for both access and disambiguation pruning is to use relative-width beam search, which prunes any construction more than a constant times worse than the best construction. If an idiom is much less probable than a lexical item, it would only be accessed if enough of the idiom had been seen to provide more evidence for the idiom that just for the lexical item. Our preliminary proposal is that there is a single fixed access beamwidth threshold for the grammar, $\alpha$, and that all constructions are accessed whose probability is within a multiple of $\alpha$ of the most likely construction. Because there is insufficient psycholinguistic data on syntactic access in particular, we do not propose a specific value for $\alpha$ in this article. We do, however, propose a specific range of values for the disambiguation beam-width in §4.

Let’s work through one example of access. Suppose we are parsing a sentence beginning ‘who can . . .’. The parser has just seen the word ‘who’, and already accessed a couple of the sentential constructions which model wh-questions, the MAIN-CLAUSE-NON-SUBJECT-WH-QUESTION and the SUBJECT-MATRIX-WH-QUESTION. Both of these constructions are marked with role [gf main] to indicate that they are main-clause-level constructions, and syn [vif fin], indicating that they will have finite verb phrases. Both of
these constructions begin with a wh-element, marked \textit{[syn \{wh +\]} in Figure 5. In the \textsc{subject-m-wh-question} this is followed by a \textsc{vp} (a fact which is inherited from the \textsc{subject-predicate} construction); in the \textsc{mc-non-subject-wh-question} this is followed by an inverted clause.

The inverted clause construction ('\textit{are you there?}', 'so \textit{will she}') requires an \textsc{aux} as its first constituent. The \textsc{vp} construction requires any verb as its first constituent. Consider now the dynamic programming chart on which the parser will keep records of interpretations in progress. The chart will have an arc for each construction being considered. Assume for simplicity that these were the only two constructions active on the chart, as shown in Figure 6. Each arc on the chart is associated with a construction, which has a dot to indicate the current place in the parse. Thus, both constructions have already seen their first constituent (marked \textit{[syn \{wh +\]}).

The probability based on top-down evidence for an \textsc{aux} appearing next given the \textsc{mc-non-subject-wh-question} is 1.0, since the construction requires an \textsc{aux} to appear. Since \textsc{aux} subsumes \textsc{can} in the type hierarchy,
the likelihood $P(\text{CAN}|\text{AUX})$ is simply $P(\text{CAN})/P(\text{AUX})$. Computing the prior probabilities of each of these construction from the Brown Corpus (Francis & Kučera, 1982), gives us $P(\text{CAN})/P(\text{AUX}) = .0017/.025 = .07$. The probability based on top–down evidence for any verb appearing next given the Subject-M-Wh-Question is 1.0, again since a verb is required. Since the prior probability of a VERB in the Brown Corpus is .085, this makes $P(\text{CAN}|\text{SUBJECT-M-WH-QUESTION}) = P(\text{CAN}|\text{VERB}) = P(\text{CAN})/P(\text{VERB}) = .0017/.085 = .021$. The probabilities of 'can' given the two Wh-Question constructions are independent, and so the actual probability based on top–down evidence of 'can' is a weighted sum:

\[(13) \quad P(\text{can}|e^+) = P(\text{MC-Non-S-Wh-Q}|\text{context})P(\text{can}|\text{MC-Non-S-Wh-Q}) + P(\text{S-M-Wh-Q}|\text{context})P(\text{can}|\text{S-M-Wh-Q})\]

$P(\text{MC-Non-Subject-Wh-Question}|\text{context})$, assuming just the context of this sentence, i.e. the first word 'who', is the probability of the derivation of this prefix 'who' from the MC-Non-Subject-Wh-Question. Whatever these probabilities are, the final $P(\text{can}|e^+)$ will be between 0.21 and .07. This is the evidence that will be combined with any bottom–up evidence for 'can'.

### 3.4 Advantages of the Model

The probabilistic model of access accounts for a number of psycholinguistic results. First, the model explains the frequency effect in lexical access. Lexical items with a higher frequency will tend to have a higher probability. We can see this by considering the equation for bottom–up evidence, applied to lexical access:

\[(14) \quad P(w \mid e) = \frac{P(e \mid w)P(w)}{P(e)}\]

(14) implies (15), that is, that the probability of any word $w$ is directly proportional to its prior probability $P(w)$, which we take from maximum likelihood estimates from relative frequencies.

\[(15) \quad P(w \mid e) \propto P(w)\]

Since idioms and lexical items are treated equivalently, the model thus also accounts for experiments on idioms which show that more familiar idioms are accessed more quickly (d'Arcais, 1993). Finally, the model predicts that this frequency effect should extend to syntactic constructions.

But the probabilistic algorithm differs from a simple frequency-based algorithm in making the prediction that access of a construction will be **inversely** proportional to the probability of the evidence. In other words, the relationship between the posterior probability of a construction and the priors of evidence and construction is as follows:
\[(16) \quad P(c \mid e) \propto \frac{P(c)}{P(e)}\]

Intuitively, things which are common are not good evidence for any one construction in particular. This aspect of using posteriors rather than priors to model access explains experimental results in both lexical and idiom processing.

First, the model provides a way to formalize the intuitions of Cacciari and Tabossi's (1988) results on idiom access. Cacciari and Tabossi distinguished predictable and unpredictable idioms in Italian. Predictable idioms are those like *in seventh heaven* in which they considered there to be an early cue to idiomaticity. Unpredictable idioms were those like *go to the devil* in which the cue that the phrase was not literal came late in the idiom. They found that the idiomatic meaning was accessed immediately for predictable idioms, but much later for unpredictable idioms (300 ms after offset). Since the authors controlled for frequency of idioms, this cannot be accounted for by assuming that more frequent idioms were accessed earlier.

But consider what it means to be a cue to idiomaticity. Essentially, the longer the structure of the idiom resembles a literal phrase, the later the cue. That is, the cue is marked by a particularly unlikely or rare word, combination of words, or construction. This is exactly what we would expect from a probabilistic approach. A more common structure, that is, one which is a very likely candidate for literal interpretation, will have a high probability. A further examination of their late-cue idioms showed that many of them also began with very common words such as *venire* ('come'), or *andare* ('go'). The probabilistic interpretation of this result, then, is that the highly probable words or constructions which began the idiom did not prove a good source of evidence for the idiom, because they provided evidence for so many other constructions as well. Again, this supports the use of a true probabilistic model, rather than the simple relative-frequency model assumed by previous lexical access theories.

Next, the probabilistic model explains the similarity neighborhood effects found by (Luce, Pisoni, & Goldfinger, 1990) and others. Luce et al. showed that words differ in their *similarity neighborhoods*, the set of words which resemble them phonetically. They showed a correlation between the time to access a word and the size of its neighborhood; words with large neighborhoods were slower on auditory naming and lexical decision. These effects are predicted by the probabilistic model just as for idiom access. The difference between a frequency-based and a probabilistic theory of lexical access are the likelihood term \(P(e|c)\) and the prior of the evidence \(P(e)\) in (14). The likelihood term will model the goodness-of-fit of a lexical entry to the input; we don't consider this further. If the evidence for a word is some (perhaps noisy) string of phones, then the evidence prior expresses how common this
string of phones is a priori, perhaps the phonotactics of the language. But consider what it means to have a large frequency-weighted neighborhood; it means that there are a lot of words which strongly resemble the target word. But if a lot of words resemble the target word, the target word must have a very common phonology. In other words, a large neighborhood means the prior probability of the evidence will be high. A small neighborhood means there are very few words in the language whose phonotactics resemble the target word, and hence the string of phones used as evidence will have a small probability.

Luce et al. (1990) explained the neighborhood effects by adding terms for neighborhood size and frequency to R.D. Luce's (1959) choice rule. We show that their modified Luce rule is an approximation to our probabilistic model. Luce et al. (1990, p. 125) gives the following equation:

$$P(ID) = \frac{P(\text{stimulus word}) \times freq_s}{P(\text{stimulus word}) \times freq_s + \sum_{j=1}^{n} p(\text{neighbor}_j) \times freq_j}$$

What they describe in their equation as $P(\text{stimulus word})$ corresponds to our likelihood term, but their method of computation blurs the distinction between likelihood and posterior. Using confusion matrices between underlying and surface phonological form, they compute $P(\text{stimulus word})$ as $P(\text{stimulus word} | \text{input string})$ rather than $P(\text{input string} | \text{stimulus word})$. For example, they compute the probability of the stimulus word /kæet/ by multiplying together the confusion probabilities $P(k|k)$, $P(æ|æ)$, and $P(t|t)$. Since with no deletions, insertions, or reductions, these probabilities are symmetric, their computation does in fact simulate the actual likelihood computation. But the actual likelihood computation would be completely different if the lexical entry for a word does not have the same string of phones as the surface string (e.g., for deletions, $P(\theta|t) \neq P(t|\theta)$).

The denominator of (17) incorporates the neighborhood size and confusability with the target word. But this equation can also be viewed as a heuristic to approximate the actual prior probability of the evidence. By definition, the true prior probability of the evidence is equal to a weighted sum over all words $l$ in the lexicon of the conditional probability of the evidence given $l$:

$$P(e) = \sum_{l \in \text{lexicon}} p(e|l)p(l)$$

What Luce et al. (1990) compute only approximates this true prior. First, they again compute the posterior $P(\text{neighbor word} | \text{input string})$ rather than $P(\text{input string} | \text{neighbor word})$. As we discussed above for the likelihood computation, the difference is not very great in their case, since they use single-phone confusion matrices and a simple string-of-phones model of phonology. But again, the more different the neighbor word is from the
target, the more likely that asymmetries in the confusion matrices will cause their equations to diverge from the true probabilities. Second, they only consider the probability of neighbor words, rather than all words in the lexicon. In practice, this is probably a good approximation, since it is the neighbor words that will provide the bulk of the probability mass, since the likelihood \( p(e|l) \) will be extremely low for words which are very different than the target word. However, it is only an approximation, and requires making an arbitrary decision about what constitutes a neighbor. Computing the true prior requires no such arbitrary definitions of neighborhood.

A final source of evidence for the Bayesian evidential theory has to do with the varying use of top-down evidence in syntax versus phonology. Tanenhaus and Lucas (1987) note that psycholinguistic evidence of top-down effects is very common in phonology, but much rarer in syntax. Researchers have noted that top-down syntactic/semantic effects on lexical access are only reported in extremely strong contexts, particularly with lexical contexts. For example, Wright and Garrett (1984) found evidence for top-down syntactic effects by showing that very strong syntactic contexts affected the reaction time for lexical decisions on nouns, verbs, and adjectives. In one experiment, a context ending in a modal verb sharply reduced the time for lexical decision to a verb. Similarly, a context ending in a preposition reduced the time for lexical decision to a noun. In phonology, the conditional probability of a phoneme appearing given a word in which it occurs is very high, and thus top-down evidence will be quite high. Syntactic constraints, on the other hand, are generally specified in terms of very abstract constructions like NOUN or VERB. Thus the top-down conditional probability of any particular noun appearing is quite low, and top-down evidence will be much lower.

The Bayesian access theory predicts that top-down syntactic effects are more likely in cases where a syntactic construction refers to a particular lexical item. Here the conditional probability of this given word occurring will be quite high. Preliminary evidence for this prediction comes from Cacciari and Tabossi (1988), who found just such top-down syntactic effects from idioms.

We can summarize the psycholinguistic data which support our probabilistic algorithm as follows:

- The access-point of a construction varies across constructions and contexts.
- Evidence for a construction is weighted in direct proportion to the frequency of the construction.
  - Lexical access time correlates directly with word frequency.
- Evidence for a construction is weighted in inverse proportion to the frequency of the evidence.
—Lexical items with high frequency-weighted neighborhoods take longer to access.
—Idioms with high-frequency words take longer to access.

- Top-down evidence for a construction is weighted in proportion to the relative specificity of the evidence and the construction.

In the next section we continue our arguments for a probabilistic model of language processing, showing that using probabilities to rank interpretations of an ambiguous utterance, just as we rank constructions to access, can account for psycholinguistic results on disambiguation.

4 DISAMBIGUATION

Natural language is inherently ambiguous. In studying the response of human subjects to this ambiguity, cognitive studies have amassed a large collection of results to be accounted for by a computational model. This includes on-line and off-line experiments on parsing preferences, (Britt, Perfetti, Garrod, & Rayner, 1992; Crain & Steedman, 1985; Ford, 1982; Frazier & Rayner, 1987; Taraban & McClelland, 1988; Whittemore, Ferraro, & Brunner, 1990) results on the interpretation of garden-path sentences, (Bever, 1970; Frazier & Rayner, 1987; Kurtzman, 1985) and studies of gap-filling/valence ambiguities (Tanenhaus, Stowe, & Carlson, 1985, etc).

Whether a model of parsing is serial or parallel, disambiguation preferences are accounted for by applying a ranking across interpretations or parse-trees. In serial parsers, the top-ranked interpretation is chosen and the others discarded. In parallel parsers, some broader set of highly-ranked interpretations is maintained.

The two classes of parsing algorithms differ more substantially in modeling garden-path effects. A serial parser models garden path effects by relying on some additional heuristics or innate parser structures which make the parser unable to build the correct interpretation. These structures might include the 3-constituent window of Marcus (1980), or the similar word-based window of the early Sausage Machine (Frazier & Fodor, 1978).

In a parallel parser, by contrast, garden path effects are usually modeled as an effect of pruning. Memory or other constraints require the parser to prune some interpretations. Using the same ranking which models disambiguation preferences, the parser prunes some set of low-ranked interpretations. The garden-path effect is explained because the correct interpretation of an utterance is among these pruned interpretations.

In this section we argue that a parallel algorithm which ranks interpretations by their probability and prunes low-probability interpretations via beam-search can explain both preference effects and garden path effects. In
addition, we show that the consequence of using probability as a metric is to choose interpretations which are the most coherent with previous expectations. §4.1 first introduces relevant data on parsing preferences and garden-path sentences, and then §4.2 presents the local coherence model of disambiguation. §4.4 will compare our model with previous models of disambiguation and garden paths, arguing that the parallel probabilistic algorithm, besides accounting for the access results discussed in §3, accounts for a much broader variety of disambiguation results than previous models.

4.1 Psychological Results
The psychological results on disambiguation may be separated into two classes. First are studies, both on-line and off-line, on human disambiguation preferences. Examining which interpretations a subject prefers (of, say, an ambiguous prepositional phrase attachment) provides data to be modeled by a potential theory of disambiguation.

For example, Ford et al. (1982), in an off-line experiment, asked subjects to perform a forced choice between two meanings of an ambiguous utterance. They showed that in (19), subjects preferred to attach the prepositional phrase on the beach to the noun dog, while in (20), subjects preferred to attach the prepositional phrase to the verb kept.

(19) The women discussed the dogs on the beach.
    a. The women discussed the dogs which were on the beach. (90%)
    b. The women discussed them (the dogs) while on the beach. (10%)

(20) The women kept the dogs on the beach.
    a. The women kept the dogs which were on the beach. (5%)
    b. The women kept them (the dogs) on the beach. (95%)

Taraban and McClelland (1988) studied a number of preposition-attachment ambiguities, by measuring preferences when subjects had seen a sentence up to and including the preposition, but not including the prepositional-object head noun. In general, they found that subjects used both verbal and nominal expectations to try to attach the prepositional objects. They found for example in (21)–(23) that subjects preferred noun phrase attachments.

(21) The executive announced the reductions in the budget / evening.
(22) The philanthropist appreciated the story on his generosity / deathbed.
(23) The high-school senior stated his goals for the future / principal.

The second kind of psychological result deals with garden-path sentences. In garden-path sentences, readers are fooled by locally ambiguous structures into selecting an incorrect interpretation. We summarize here a number of cases of ambiguities that lead to the garden-path effect as well as some that do not.  

* We follow Gibson (1991) in marking garden-path sentences with a pound-sign (#).
1. Main Clause-Relative Clause ambiguity is often resolved in favor of the main clause analysis.

(24) # The horse raced past the barn fell. (Bever, 1970)

However, semantic context can reduce or eliminate the garden-path effect (Crain and Steedman, 1985).

(25) # The teachers taught by the Berlitz method passed the test.
(26) ?The children taught by the Berlitz method passed the test.

2. Lexical category ambiguities sometimes cause garden-path effects, sometimes do not.

(27) # The complex houses married and single students and their families. (Hearst, personal communication, December, 1991)
(28) The warehouse fires destroyed all the buildings. (Frazier & Rayner, 1987)
(29) The warehouse fires a dozen employees each year. (Frazier & Rayner, 1987)
(30) # The prime number few. (Milne, 1982)
(31) # The old man the boats.
(32) # The grappling hooks on to the enemy ship. (Milne, 1982)

3. Valence ambiguities sometimes cause garden path effects

(33) # The landlord painted all the walls with cracks.
(34) # Ross baked the cake in the freezer. (from Hirst (1986))

Any cognitive theory of disambiguation must explain why certain ambiguities cause processing difficulties and why others do not.

Indeed, recent psycholinguistic results have been used to argue that this distinction between disambiguation preferences and garden-path sentences is only an approximation, and that processing difficulty is better modeled as a continuum, with slight preferences at one end and garden-path-like pruning at the other (MacDonald, Pearlmuter, & Seidenberg, 1994; Tabossi, Spivey-Knowlton, McRae, & Tanenhaus, 1994). Since a probabilistic algorithm maps sentences onto the real numbers, it is well-suited to modeling this kind of continuum. In the absence of sufficient data on the details of the processing difficulty continuum, however, the rest of this section will simply show that our coherence-based algorithm models the parse preference effects, and that the same preference ranking together with a pruning algorithm can be used to predict the garden-path effect for the garden-path data.

4.2 The Model

§3 showed that a parallel algorithm which uses probabilities to rank constructions could be used to model lexical and constructional access. This section shows that the same probabilistic ranking can be used to account for disambiguation preferences. Each interpretation will be ranked by its probability, computed from both constituent and valence probabilities. We
show this ranking is able to account for the psychological data on disambiguation preferences.

A disambiguation model must also explain the processing difficulties that garden-path sentences cause human parsers. The serial parsing models that were traditional in psycholinguistics model garden path effects by relying on some particular heuristics or innate parser structures which make the parser unable to build the correct interpretation. For example, because of its 3-constituent window, Parsifal (Marcus, 1980) is unable to simultaneously view the initial noun phrase 'the horse' and the final word 'fell' in (24). Since the parser is required to build structure before 'the horse' moves out of its window, and since there is no evidence yet for the reduced-relative parse, Parsifal will build the main-verb parse. Thus a hardwired memory constraint is used to force the parser to drop the correct parse from consideration.

In a parallel parsing model, however, there is a natural alternative to these specific heuristics or parser structures, proposed by Kurtzman (1985), Norvig (1988), and Gibson (1991). In their models, garden-path sentences are explained by the same ranking that is used to model disambiguation preferences. The parser ranks each hypothesis, and then relies on some sort of constraint (such as limited-memory) to force low-ranked hypotheses to be pruned. The garden-path effect is explained because the correct interpretation of an utterance is among these pruned hypotheses.

In our model, the parser ranks each parse by its probability, and uses beam-search to limit the search-space. As Gibson (1991) proposes, we can find an empirically acceptable beam width and also test our model by considering the probabilities of each interpretation of ambiguous sentences. A sentence where there is evidence for pruning (such as the garden-path effect) sets an upper bound on the beam width. A sentence where there is no evidence for pruning (both interpretations are acceptable) sets a lower bound on the beam width. We show that a beam-width ratio of anywhere between 3.8:1 and 5.6:1 between the best and the pruned hypothesis is sufficient to account for a number of garden-path examples. Table 1 summarizes some of the garden-path/embedded anomaly and nongarden path examples which we will work through in the rest of this section. In the first three examples (explained in further detail here; †indicates a local embedded anomaly) the ratio between the probabilities of the best and pruned hypothesis is 5.6 or

| The complex houses married and single students and their families | 267:1 |
| The sheriff wasn't sure which rock the cowboy raced... (desperately past) | 12:1 |
| The district attorney found out which church the reporter asked... (anxiously about) | 5.6:1 |
| The warehouse fires destroyed all the buildings | 3.8:1 |
| The bird found in the room died. | 3.7:1 |
greater. In the last two examples, where the local ambiguity does not cause a garden path, the ratio is 3.8 or less.

We turn now to the actual computation. §2 summarized the two kinds of probabilities that augment constructions, constituent probabilities and valence probabilities. The constituent probability is the standard SCFG rule-probability, and as we discussed an SCFG assumed that the probability of each rule or construction is independent, and hence that the probability of a parse tree can be computed by multiplying the probabilities of the individual rules. We also make the simplifying assumption that valence probabilities are independent of constituent probabilities and of each other. Thus, the likelihood of a sentence given an interpretation is the product of each constituent and valence probability in its derivation. Figure 7 shows the probabilities associated with a small noun phrase beginning a sentence.

4.3 Modeling Preference Effects

Lexical Valence Preference
Consider the Ford et al. (1982) results on attachment preferences for keep and discuss summarized above. Ford et al. explained their results by proposing that keep and discuss each had two possible subcategorization frames; one with a single <NP> complement, one with <NP PP> . They proposed that the verbs differed, however, in their preferences for these valence structures; keep prefers <NP PP>, while discuss prefers <NP>.

Our interpretive model is conveniently set up to test Ford et al.'s intuitions, since both constructions and valence structures are annotated with conditional probabilities. We assume that keep and discuss each have two valence frames. For discuss, we assume for simplicity that the two frames are defined strictly syntactically as <NP> and <NP PP>. For keep the frames are somewhat more complex to define. While the optional second argument of keep can be a PP (keep the dogs on the beach), Table 2 shows that it can also be an AP, a VP, or a particle.
TABLE 2
Complements of *keep*

| AP | keep the prices reasonable |
| VP | keep his foes guessing    |
| VP | keep their eyes peeled    |
| PRT| keep the people in        |
| PP | keep a number of surprises under our hats |
| PP | keep his nerves from jangling |

TABLE 3
Valence Probabilities Computed from the Penn Treebank

| discuss | <NP PP> | .24 |
|         | <NP>    | .76 |
| keep    | <NP XP[pred +]> | .81 |
|         | <NP>    | .19 |

Most theories model this class of complements syntactically, as *secondary predicates*, represented as *[pred +]*, although it is also possible to capture them thematically as *RESULTS*. Assuming, again only for simplicity, the *[pred +]* representation, *keep* has the thematic grid options <NP> and <NP XP[pred +]>

We used the Penn Treebank to compute probabilities for each valence structure. The results are shown in Table 3.

If we only consider these probabilities, we arrive at a result consistent with the Ford et al. (1982) data. Since *discuss* prefers a lone NP complement, the PP will attach to the noun phrase, while *keep* prefers to attach the PP as an argument of the verb.

However, the different attachments of the preposition phrases also lead to different phrase structures. Figure 8 shows the different phrase structures for each interpretation. The interpretation in which *on the beach* is a complement of 'keep', is shown. This interpretation includes the VP-level rule with a slot for oblique complements:

(35) [.15] VP — V NP XP

The probability of this interpretation includes this probability, .15, as well as the valence probability that *keep* fills its RESULT argument, or .81. All the other probabilities involved in the interpretation are shared with the other derivation, and so we need not consider them when comparing the two. The second interpretation has a slightly more complicated phrase structure at the NP level, and so there are two rules in this parse that are not in the other:

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*We assume here a GPSG-like treatment of VP rules; generalizations across different VP types are captured by an abstract VP in the type hierarchy.*
(36) [.39] VP → V NP
(37) [.14] NP → NP Postmodifier

Note that the probability of the simple transitive VP is higher than the probability of the multiple-complement VP above. This is offset, however, by the lower probability assigned to the unfilled RESULT valence slot, and in addition by the lowered probability due to the extra NP rule. Thus the probability of the verb attachment is .12/.01 = 12 times the probability of the noun attachment, roughly modeling the Ford et al. (1982) results, in which the verb attachment was preferred 19 times over the noun attachment (95%).

Figure 9 shows the two interpretations of *discuss the dogs on the beach*. Although the syntactic structures are the same, the different valence probabilities of *discuss* and *keep* will cause the model to prefer the opposite interpretation. Although the prediction is correct, the model's 53% preference for the NP attachment does not quantitatively match the 90% preferences in the Ford et al. (1982) results. One possible explanation for this is the forced-choice nature of the Ford et al. experiment. Requiring the subjects to choose between the two interpretations may have caused the experiment to overestimate the probability of the most-preferred interpretation.

Augmenting valence slots with probabilities also allows us to explain with no further assumptions the stronger preferences for obligatory over optimal arguments. Obligatory arguments will have unity or near-unity probabilities, while optional arguments will have lower probabilities. Britt (1991) showed, for example, that an obligatory argument slot placed a
stronger expectation to be filled than an optional one. In an on-line experiment, Britt showed that when choosing between noun and verb attachment of ambiguous preposition phrases in two-referent contexts, subjects preferred to attach the PP to the verb if the verb obligatorily subcategorized for the PP. If the verb only had an optional subcategorization for the PP, there was a preference for noun attachment instead.

Viewing the probabilistic model in a more general way, we can explain the general preference for arguments over adjuncts as resulting from a combination of valence probabilities and constituent probabilities. The arguments will have a nonzero valence probability, while adjuncts do not fill a valence slot in a predicate, and so not attaching an adjunct incurs no probabilistic penalty. This is true for both nominal and verbal adjuncts. In addition, when choosing between verbal arguments and nominal adjuncts, since nominal adjuncts will require an extra stochastic context-free rule to adjoin with, they will have a lower probability than a simpler structure that has no such rule. This is because due to the independence assumption inherent to context-free grammars, the probabilities of each rule are multiplied together. Thus a more complex structure will have more probabilities to multiply, and since each probability is less than 1, a lower total probability. Note that the fact that a more complex structure is likely to have a lower probability than

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This is especially true if we assume the common model in which complements are the arguments of verbs but adjuncts act as predicates in their own right.
a simpler structure may explain the success of heuristics like Minimal Attachment, which preferred attachments producing the simplest syntactic structure. However, a simpler structure will not necessarily have a higher probability, since it is not just the cardinality of the probabilities but their values which matter.

In order to explain this sort of preference, Ford et al. (1982) and Abney (1989) proposed a heuristic that preferred attachments to verbs over attachments to nouns. Relying on valence probabilities eliminates the need for this occasionally faulty heuristic.

4.4 Modeling Garden Path Effects
In our model, as in most explanations for the garden-path effect, the effect arises when the processor chooses an interpretation which fulfills local linguistic expectations over one which is globally plausible, pruning the more plausible interpretation. Testing this model requires showing that the theory predicts pruning in a case of ambiguity whenever the garden-path effect can be shown to occur. If the beam width is set too wide, the theory will mislabel a garden path sentence as merely a less preferred interpretation. Conversely, if the window is too narrow, the theory will mislabel parseable sentences as garden-paths.

In many ways we view our model as an extension of the parallel model of Gibson (1991). Gibson first showed that a parallel (although nonprobabilistic) model could account for the wide range of syntactic garden-path effects, and proposed that probabilistic garden-path effects might also be modeled with a parallel algorithm. In the rest of this section, we work through some examples of garden paths, showing that the coherence model can not only account for garden paths that seem obviously probabilistic in nature (such as (27) and (31) above), but also more traditional garden-path sentences such as (24). We consider these garden paths in three classes: those caused by constituent probabilities, those caused by valence probabilities, and those caused by combinations of the two.

We will show that a beam-search in which interpretations are pruned if they have less than about 1/5 the probability of the best interpretations can account for all of the data we consider.

4.4.1 Garden Paths Caused by Construction Probabilities
For simplicity, we begin by considering garden paths caused by simple phrase-structure probabilities. Consider the garden path sentence (38). In the intended interpretation of the sentence, 'complex' is a noun, and 'houses' is a verb; thus the students are housed by the complex. However, most readers initially interpret 'the complex houses' as a noun phrase, and are confused by the lack of a verb.
(38) # The complex houses married and single students and their families. (Hearst, 1991)

Figure 10 shows the two interpretations of the prefix the complex houses. For convenience (and in order to make the model maximally general) we have shown the phrase structures in a simplified minimal-assumption framework.

In the initially preferred interpretation in (a), 'houses' is the head noun and 'complex' is an adjective. In the dispreferred interpretation in (b), 'houses' is a verb and 'complex' the head noun. The difference in probability between the two interpretations is due mainly to the different lexical category probabilities. (39)–(44) show conditional probabilities generated from the Brown Corpus. Notice that 'houses' is more likely to be a noun than a verb, while 'complex' is more likely to be an adjective than a noun. In addition, we have shown the probabilities of the plural and 3rd-person-singular interpretations of the ambiguous 's' morpheme.

(39) [.0024] N → house
(40) [.0006] V → house
(41) [.00086] Adj → complex
(42) [.000029] N → complex
(43) [.23] N → ROOT s
(44) [.086] V → ROOT s

The nonlexical nodes of the parse tree seem irrelevant here because the rest of the trees differ in probability only by about 20%. Computing the complete probabilities of the two interpretations correctly predicts the garden path effect; first, the probabilities correctly prefer the (a) interpretation, and second this reading is 267 times more probable than the dispreferred reading. Since this is much greater than the beam width, interpretation (b) will be pruned.

By contrast, (45) shows an example due to Frazier and Rayner (1987) where a similarly ambiguous construction does not cause the garden-path effect.

(45) a. The warehouse fires destroyed all the buildings.
    b. The warehouse fires a dozen employees each year.
Note that the difference between (38) and (45) cannot be due to valence differences or nonprobabilistic structural differences, since the two examples have very similar valence and syntactic structure. The difference between the two examples is in the lexical probabilities.

The difference in probability of lexical category for fire is shown in (46)-(47). The two interpretations in Figure 11 have probabilities of $4.2 \times 10^{-5}$ and $1.1 \times 10^{-5}$; their ratio of $3.8/1$ is less than the beam-width, and hence neither interpretation will be pruned.

(46) [.00072] $N \rightarrow fire$

(47) [.00042] $V \rightarrow fire$

4.4.2 Garden Paths Caused by Valence Probabilities

For an example of the garden path effect caused by valence probabilities, we turn to an experiment which was originally designed to test Fodor's (1978) Lexical Expectation Model of gap-finding in a serial parser. Fodor was attempting to model how the processor decided which gaps to fill with a given fronted element. For example, the fronted NP the book in (48a) could turn out to be the object of want or the object of some embedded verb like buy.

(48) a. I saw the book you wanted . . .

b. Book is object of want: I saw the book you wanted.

c. Book is object of deeper verb: I saw the book you wanted me to buy.

Fodor proposed that the human sentence processor posits gaps only following verbs which are more likely to be transitive. That is, given a fronted argument, when the parser reaches a verb it should immediately attach the argument as the direct object of the verb if the verb is more likely to be transitive, otherwise it should wait. Assuming that ask is preferentially transitive and race preferentially intransitive, the parser would attach 'the man' as a direct object in the man you asked, but wait in the book you raced (assuming rather that 'the book' should be attached as the object of some lower predicate as in the book you raced past). Tanenhaus et al. (1985) tested this idea by creating test sentences with fronted arguments which make sense
if the whole sentence is seen, but which are semantically anomalous direct objects of the verb. Thus, if the parser tries to fit them into the verb right away, the reader should have a minor processing breakdown. If Fodor's (1978) model is correct, the parser should suffer an anomaly when trying to fit a semantically implausible filler into preferentially transitive verbs like ask, but should not break down on preferentially intransitive verbs like race. The experiment used sentences like 49–50, testing for an anomaly effect at the verb.

(49)  
a. The district attorney found out which witness/church the reporter asked—about the meeting. (*early gap*)
b. The district attorney found out which witness/church the reporter asked anxiously about—. (*late gap*)

(50)  
a. The sheriff wasn't sure which horse/rock the cowboy raced—down the hill. (*early gap*)
b. The sheriff wasn't sure which horse/rock the cowboy raced desperately past—. (*late gap*)

The results of the experiment supported Fodor's (1978) model; for transitive-preference verbs, there was an anomaly effect at the verb, while for intransitive preference verbs, there was no anomaly effect at the verb. Tanenhaus et al.'s (1985) explanation was based on a serial architecture; for verbs which are preferably transitive, the parser hypothesizes a gap; for verbs which are preferably intransitive, the parser does not hypothesize a gap.

We argue that rather than requiring a special gap-hypothesizing mechanism, that these results can be accounted for with our parallel architecture without proposing any special mechanism. First, as with the other valence ambiguities we have seen, the parser will activate both possible valence interpretations for each verb, one transitive and one intransitive. However, in each case the low-probability interpretation will be pruned. Because the low-probability intransitive interpretation of ask will be pruned, the parser will attempt to bind church as the direct object of ask, causing an anomaly. Because the low-probability transitive interpretation of race will be pruned, the parser will not attempt to bind rock as the direct object of race, avoiding the anomaly.

Figure 12 and Figure 13 show the disambiguations for race and ask respectively. The valence probabilities were determined from Connine et al. (1984). For race the ratio of the two interpretations is .92/.08 = 12/1. For ask the ratio is .79/.14 = 5.6/1. In both cases this is sufficient to cause pruning of the dispreferred interpretation.

4.4.3 More on Horse-Racing: Combinations of Valence and Construction Probabilities

We have now seen examples of garden-path sentences caused by construction probability differences, and by valence probability differences. In this sec-
tion we argue that Bever's (1970) familiar garden-path sentence in (51) is caused by a combination of both valence and construction probabilities.

(51) # The horse raced past the barn fell.

Here the phrase *raced past the barn* is ambiguous between a reduced relative clause VP and a main-clause VP. Clearly the garden path effect arises because the main verb reading of *race* is somehow preferred by the parser, while the reduced-relative clause turns out to be correct one.

We show first that the preference for the main-verb reading arises from our probabilistic model, and second that the difference between the two interpretations is sufficient to cause the reduced-relative interpretation to be pruned.

Figure 14 shows the two relevant interpretations of the sentence prefix *the horse raced*. Figure 14(a) shows the main verb interpretation. Note that the intransitive valence of *race* has a probability of .92.⁹

The reduced relative interpretation requires the transitive valence of *race*, since it requires the first NP to be bound as an object of the verb. In addition, this interpretation has a more complex syntactic structure; note that in addition to the syntactic rules used in the trees in (a), the tree in (b) has an additional rule:

(52) [.14] NP → NP XP

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⁹ These numbers are from Connine et al. (1984).
The combination of the extra (and low-probability) rule and the lower valence structure makes this interpretation 82 times less probable than the best-ranked main-verb interpretation. Since this places it outside the 5x beam width, the reduced-relative interpretation will be pruned.

Early attempts to explain the garden-path effect of this sentence relied solely on its syntactic structure. Bever's (1970) original proposal relied on a heuristic to prefer main verbs over reduced verbs, Marcus's (1980) relied on the fact that a 3-constituent window would not include the final verb, and heuristics like Minimal Attachment (Frazier & Fodor, 1978) rely specifically on the parse tree.

However, Pritchett (1988) recently showed that solely constituent-based solutions to the problem are insufficient, as sentences like (53) do not cause the garden-path effect, despite having exactly the same syntactic structure as (51).

(53) The bird found in the room died.

The difference between (53) and (51) is the valence structure of the verbs. Where *race* is preferably intransitive, *found* is preferably transitive. Figure 15 shows the three interpretations of the prefix "the bird found".

Again, the two main-verb interpretations are too close to each other to cause pruning. Significantly, however, the difference between the main and reduced-relative interpretations is much less than it was for *race*, the probability ratio being \( (0.62 \times 0.48)/ (0.62 \times 0.14 \times 0.92) \) or 3.7/1. This ratio is less than the beam width, and so the reduced relative interpretation will not be pruned.

4.5 Recent Theories of Disambiguation

If every theory of disambiguation relies at its core on some metric for comparing interpretations, theories may be classified by the kind of metrics they employ. Many models use *global metrics*, in which each structure is ranked by a single global criterion, such as our ranking based on conditional probability, or such as a preference for the most semantically plausible interpreta-
such global metrics are particularly commonly used by the interactive or constraint-based models of sentence processing such as the model presented here as well as others like MacDonald (1993), McClelland et al. (1989), Spivey-Knowlton and Sedivy (1995), and Trueswell and Tanenhaus (1994). In contrast, local structural models choose between interpretations by using local, domain-specific heuristics based on simple structural patterns, such as choosing local attachments, or choosing syntactically simpler structures. These local metrics are commonly employed in the two-stage parsing models, in which a modular syntactic processor makes initial disambiguation decisions based only on local syntactic information, without access to lexical and semantic knowledge. These decisions may then be reanalyzed by a second, knowledge-rich stage of processing.

This section first surveys two classes of global heuristics, those based on semantic plausibility and those based on syntactic or textual coherence. We will argue that the semantic plausibility model of Crain and Steedman’s (1985) and others cannot account for the full range of experimental results on garden-path effects, while Pritchett’s (1993) coherence-based model is insufficiently general to account for lexical access effects and frequency effects. Finally, we compare our probabilistic algorithm to other parallel algorithms (Gibson, 1991; Spivey-Knowlton, 1994). We next consider the local heuristics; as §4.3 suggested, a preference for minimal structures, all else being equal, falls out of the independence assumptions inherent to context-free grammars. We focus in this section, therefore, on the locality heuristics, arguing that, although the human parsing clearly displays some effects of locality, no previously proposed locality heuristic is sufficient to account for the range of data.
4.5.1 Global Metrics
The most well-known global heuristic for modeling human parsing preferences and garden-path data, and one that has perhaps the most intuitive appeal, is simply to choose the most plausible interpretation. This model was expressed most familiarly by Crain and Steedman (1985, p. 330):

_The Principle of A Priori Plausibility._ If a reading is more plausible in terms either of general knowledge about the world, or of specific knowledge about the universe of discourse, then, other things being equal, it will be favored over one that is not.

A number of researchers have proposed models along these lines, including Kurtzman (1985), Altmann and Steedman (1988), and Charniak and Goldman (1988). One disadvantage of this model is that in most cases it is hard to see how to make it operational. For example, what is it about having a discussion about dogs which are on the beach that makes it more plausible than having a discussion on the beach about dogs? The model also has serious difficulties in accounting for garden-path sentences. For example, Crain and Steedman (1985) argue that the preference for the main clause interpretation in main-clause/reduced-relative ambiguities is a corollary of plausibility. This because the reduced relative clause reading requires more presuppositions than the main clause reading (in particular the presupposition in ‘the horse raced past the barn fell’ that there are multiple horses). However, this model incorrectly predicts the garden path effect for (54):

(54) The bird found in the room died.

The problem is that the plausibility model does not take into account verbal valance preferences. Finally, the model cannot account for the frequency effects summarized previously. This is not to deny the role of plausibility in parsing data; as §5 suggests, one advantage of our probabilistic model is that in principle it allows the inclusion of plausibility information along with any other sources of evidence in ranking interpretations.

Most other global or global-like models use something like coherence as the metric for ranking interpretations. The roots of coherence as a processing heuristic lie in Wilks' (1975a, 1975b) rediscovery of the Joos Law (Joos, 1972), which argued for choosing a meaning which was most redundant and hence most coherent with the context (see also Hill, 1970; Joos, 1958). The idea was then taken up by the coherence-based marker-passing algorithms (Hirst, 1986; Hirst & Charniak, 1982; Norvig, 1987). More recently, Ford et al. (1982) added the insight that verbal valence was the key factor in modeling parsing preferences and garden-path sentences, incorporating a preference for argument attachments over adjunct attachments, and also proposed preference augmentations to valence.

The modern models which account for the most data are Gibson (1991) and Pritchett (1993). Both of these models, like our probabilistic approach,
emphasize coherence as a parsing heuristic. Pritchett's model is based on a serial parser which often reanalyzes its previously-produced parse, but which garden-paths in certain kinds of reanalyses:

**Theta-Reanalysis Constraint:** Syntactic reanalysis that reinterprets a θ-marked constituent as outside of current θ-domain is costly.

When faced with an ambiguity, Pritchett's parser chooses the most θ-coherent parse; i.e., the one in which the most θ-grid slots are filled, and the most arguments have θ-roles. The model accounts for 'the horse raced past the barn fell', because *the horse* is originally parsed as the subject of *race*, and hence within *race*'s θ-domain. When the parser sees *fell*, it will need to reanalyze *the horse* as within the θ-domain of *fell*, causing the garden-path effect.

Although Pritchett's model accounts for a wide variety of syntactic garden-path effects, it is unclear how it would be extended naturally to account for the probabilistic effects in (38), or the semantic or context effects in (25)–(26). In addition, of course, Pritchett's model requires disambiguation to be completely independent of access and its frequency effects, despite the evidence that lexical disambiguation at least is strongly affected by access frequencies.

Gibson (1991) proposed a parallel parsing model in which each parse of an ambiguous input was ranked by its *processing* load. Interpretation proceeded by beam search; parses were pruned if their load was significantly higher than the best interpretation. Load was defined with three separate load parameters:

1. **Property of Thematic Reception:** parses have processing load for each argument lacking a θ-role.
2. **Lexical Requirement:** parses have processing loads for each θ-grid slot lacking a filler.
3. **Thematic Transmission:** parses have processing load if a θ-role is passed through a chain (i.e., wh-movement or comp).

Like Gibson's model, our probabilistic model will still disprefer interpretations if they have unfilled slots or unparsed arguments. In a generalization over Gibson's model, however, our model can distinguish between two interpretations both of which have valence slots filled, but with different probabilities. In addition, our model can explain why the parser will sometimes prefer not to fill a slot with an argument, as in (55); as we argued in Figure 12, in this sentence because 'race' is preferably intransitive, the NP 'which rock' does not fill the direct object slot in 'race'.

(55) The sheriff wasn't sure which rock the cowboy raced...

Our explanation of reduced-relative/main-clause ambiguities is also different from Gibson's, although both models rely on the valence of the main
verb to account for *race*/*found* distinctions. Where Gibson appeals to the Property of Thematic Transmission to account for part of the difficulty of 'The horse raced past the barn fell', our model appeals to the low prior probability of the reduced relative clauses. The advantage of a probabilistic account, again, is that it explains the garden-path examples based on frequency with the same mechanism as the structural garden-paths. Although Gibson points out that his parallel model could easily be extended to deal with frequency effects, it would presumably require adding another property or set of properties; our theory accounts for the data with a single model.

In our model, as we've described it so far, as well as in the Gibson (1991) and Pritchett (1993) architectures, processing difficulty is a binary effect of parse pruning; sentences are either difficult to process or not. Because of this, the model is unable to account for processing difficulty variation within garden-path sentences (MacDonald et al., 1994; Tabossi et al., 1994). The probabilistic competition model of Spivey-Knowlton (1994) and Spivey-Knowlton and Tanenhaus (1995) attempts to address this lack by modeling processing difficulty as a continuum, caused by the time-course of competition between syntactic alternatives. Spivey-Knowlton and Tanenhaus (1995) combine information from discourse (in terms of the discourse context probability of a reduced relative), from verb frequency information (the probability of a verb appearing in the simple past tense form versus the past particle form), and from parafocal information (the presence of "by" after the verb). They estimate probabilities for each of these information sources, and then combine them (via recurrent feedback and normalization) on a set of sentences with reduced-relative/main-clause ambiguities. They show that the durations of competition between the syntactic alternatives predicts the processing difficulty in first pass reading times at the verb.

Since our probabilistic metric already maps each possible interpretation of an ambiguous sentence onto a continuous space of probabilities, we hope to investigate whether our algorithm could incorporate the insights of competition-based approaches like that of Spivey-Knowlton and Tanenhaus (1995) or Stevenson (1993) in order to model the continuum of processing difficulty.

4.5.2 Structural Heuristics

Structural heuristics were originally proposed by Kimball (1973) to explain the ungrammaticality of certain sentences, or the inability of subjects to get some readings of an ambiguous sentence. Since then a number of heuristics have been proposed, falling into two classes: heuristics to build the syntactically simplest structure (*minimality heuristics*) and those to combine nearby structures (*locality heuristics*).

However, many authors (Gibson, 1991; Kurtzman, 1985; Norvig, 1988; Osterhout & Swinney, 1989; Schubert, 1986) have noted that it is quite easy
to choose particular lexical items or particular contexts which reverse any of
the heuristics, particularly the minimality heuristics. Thus, more recent
models incorporate local heuristics in a more sophisticated way; as the first
stage in a two-stageparser. This first stage is an encapsulated module which
makes a very early, local decision about structure. This decision is reviewed,
and possibly reanalyzed, by a second stage which has access to richer knowl-
edge sources. A number of papers have argued that only minimality and
locality are relevant to the first-stage process, and that garden-path data
and grammaticality judgements have no access to this early stage of process-
ing (Clifton, 1993; Clifton, Speer, & Abney, 1991; Ferreira & Clifton, 1986;
Mitchell, 1989). Trueswell and Tanenhaus (1994), in contrast, have argued
that results such as those of Ferreira and Clifton (1986) can be reinterpreted
to argue for a more constraint-based view of processing.

A resolution of these experimental differences awaits further experiment-
al work; however, in the rest of this section we argue that formulating any
single locality heuristic which explains even the available experimental data
will be quite difficult. One major weakness of minimality heuristics, for ex-
ample, is that they are dependent on quite particular assumptions about the
grammar. As Norvig (1988) points out, for example, a syntactic simplicity
heuristic like Minimal Attachment makes no sense in a categorial grammar,
in which every derivation has the same number of nodes. In addition, since
the local heuristics only account for garden-paths and preferences due to
syntactic effects, a two-stage theory would require other, presumably un-
related mechanisms to account for garden paths due to frequency effects or
to nonlocal and semantic expectations. An interactive or constraint-based
approach to disambiguation accounts for both syntactic and nonsyntactic
effects with a single mechanism.

Another problem with the use of local heuristics and the serial disambig-
uation strategies that usually accompany them is the well-accepted evidence
that disambiguation is not always immediate. Frazier and Rayner (1987),
for example, showed that the parser was able to delay disambiguation in
(45) above (repeated in (56) below), avoiding the garden path.

(56) a. The warehouse fires destroyed all the buildings.
b. The warehouse fires a dozen employees each year.

The only way to model these effects in a serial parser is to add a complex
delay function, for which there is no independent motivation. In a beam-
search parallel parser, on the other hand, delay is simply caused by the fact
that the two interpretations in (56) have probabilities within the beam. The
existence of lexical ambiguities like (56) which do not cause the garden-path
effect also argues against Ford et al.'s (1982) proposal that it is lexical
category reanalysis which explains the difficulty of recovering from garden
paths.
In the remainder of this section we present a paradox for locality heuristics. We show that there are two styles of locality heuristic, and that experimental results and grammaticality judgements which have been taken as evidence for one kind of locality heuristic cannot be modeled by the other. We are not arguing that locality plays no role in language processing; it clearly does. Besides the experimental results discussed above, for example, Gibson and Pearlmutter (1994) found evidence for locality heuristics in a statistical analysis of preposition-phrase attachment in the Brown corpus. However, we argue that although each version of locality accounts for some data, there are principled problems with generalizing any of the locality models to handle all the data. Thus, the intuition that locality plays some role in processing is very difficult to make operational.

Locality principles described in the literature include Right Association (Kimball, 1973), Local Association (Frazier & Fodor, 1978), Late Closure (Frazier, 1978), Final Arguments (Ford et al., 1982), and Graded Distance Effect (Schubert, 1984, 1986), Rule B (Wilks, Huang, & Fass, 1985), Attach Low and Parallel (Hobbs & Bear, 1990), and Recency Preference (Gibson, 1991). These models of locality fell into two classes. In the windowing models, locality effects are explained by appealing to a limited buffer or window on the input sentence, generally assumed to model some sort of limited memory. Since the parser is limited to working within this window, constituents will necessarily be attached locally if they can be. The window may be limited by the number of words (Frazier & Fodor, 1978; Magerman & Marcus, 1991) or the number of constituents (Marcus, 1980).

In the iterative or right-to-left models, when attempting to attach a constituent, the parser moves right to left until it finds the first appropriate head to attach the constituent to. Unlike the windowing models, in these models locality does not fall out of the structure of the parser, and must be stipulated. Iterative models differ in how they define the notion of appropriate head—this may be semantically appropriate (Wilks et al., 1985; Hobbs & Bear, 1990; Whittemore et al., 1990) or syntactically appropriate (Gibson, 1991).

For example, the Local Association principle of Frazier & Fodor (1978), is a windowing model. It establishes a fixed-length buffer which can hold five or six words, and predicts that locality effects can be explained by the limited view imposed by this buffer. While windowing models have the advantage of explaining locality rather than stipulating it, it has proven to be very difficult to satisfactorily define a window size, as we will see below.

The Recency Preference principle of Gibson (1991) is an iterative or right-to-left model. In this model, whenever there is more than one possible attachment point for an adverbial, all but the most recent attachment point are removed from consideration. In his case, adverbials can be attached to verbs or sentences, so the Recency Preference principle requires that an adverb cannot “skip” a local verb and attach to a more distant one.
Another model, again an iterative one, assumes what might be called the *Most Recent Semantically Compatible Attachment* principle, first proposed by Wilks et al. (1985) (as the Rule B algorithm of the CASSEX program), and used also by Hobbs and Bear (1990) and Whittemore et al. (1990). These models choose among attachments by attempting to attach an adverbial to each possible head, starting with the most recent, and moving further left, and selecting the first one that fits semantically.

We discuss the performance of these models on three classes of data: garden-path sentences, restrictive relative clause attachment, and verb-particle attachment.

A number of parsing models have attempted to use windowing locality to explain the garden-path effect (although this has not been proposed with iterative models). Consider (57):

\[(57)\]  
\#
The horse raced past the barn fell

All windowing models predict the difficulty of this sentence by arranging that only *The horse raced past the barn* is in the window. In the Marcus (1980) parser, the three constituent buffer is filled by *the horse*, *raced*, and *past the barn*; in the Local Association model the buffer is exactly six words long.

The problem with the windowing models, as has been noted by Norvig (1988), Gibson (1991), Schubert (1986), and Pritchett (1993), is that for any given window size, it has been shown possible either to construct a parsable sentence which does not fit inside the window, or an unparsable sentence which does. For example (58) fits entirely inside the 6-word Local Association buffer, and hence the model incorrectly predicts it to be processable. Conversely, Marcus' 3-constituent window incorrectly predicts that (59) is a garden-path sentence.

\[(58)\]  
\#
The horse raced yesterday fell

\[(59)\]  
I know my aunt from Peoria died

Since these kinds of arguments have been made before, we turn to a new analysis of some non-garden path data. It has been argued that locality heuristics can account for the three kinds of preferences and grammaticality judgements in (61)–(62). First, Kimball (1973), Wanner (1980), and Gibson (1991) argue that iterative locality is what causes the adverb ‘yesterday’ to attach to the local verb ‘died’ rather than the distant verb ‘said’ or ‘thought’ in (60). Second, Kimball also argued that locality explained why (61a) could only mean that the job was attractive, not that the woman was attractive, and thus could not have the same sense as (61b). Finally, Kimball and Gibson have argued that iterative locality explains why examples like (62) are unacceptable, since the particle ‘out’ prefers to attach to the nearer predicate ‘take’ rather than ‘figure’.
(60)  a. Bill said John died yesterday.
     b. Bill thought John died yesterday.

(61)  a. The woman took the job that was attractive.
     b. The woman that was attractive took the job.

(62)  Joe figured that Susan wanted to take the train to New York out.

We argue that (61) cannot be explained by locality, and that no single locality heuristic can explain both (60) and (62). Consider the processing of restrictive relative clauses. Before Kimball's work, it was generally assumed that sentences such as (63a,b) (from Hankamer, 1973) were simply ungrammatical:

(63)  a. *A man whom I married my sister who had castrated himself.
     b. *I gave a kid a banana who was standing there looking hungry.

Kimball claimed that sentences like (63a,b) must be grammatical, because they were created by the same Extraposition from NP transformation that created (64b) from (64a). Since (Kimball claimed) (64b) was grammatical, (63a,b) must also be grammatical, and must only be ruled out for performance reasons. Thus the principle of Right Association would attach the phrase who had castrated himself to the noun sister instead of man in (63b), and thus (63) would be grammatical but unparseable.

(64)  a. The woman that was attractive fell down.
     b. The women fell down that was attractive. (grammatical according to Kimball's theory)

Our informants universally agreed, however, (although with no attempt at experimental confirmation) that (64b) is not at all grammatical, and so we star it for future reference.

(65)  *The woman fell down that was attractive.

But (65) cannot be unacceptable because of either windowing locality or iterative locality. It cannot be accounted for by windowing locality since there are only two words (and only one constituent) between the relative clause and the nominal head, and no windowing theories predict such a small window. It cannot be unacceptable due to iterative locality, since there is no intervening nominal head; indeed there is no other nominal head in the sentence at all!¹⁰

The next phenomenon which is commonly cited as evidence for locality principles is the attachment of verbal particles to their head verbs. Kimball (1973) first presented the following examples, arguing that locality explains

¹⁰ One reviewer pointed out that making the relative clause heavier makes this sentence grammatical: "The woman fell down who had just stepped onto the moving platform." However, heavier phrases can often appear in sentence-final position in cases where nonheavy constituents are disallowed, and thus the acceptability of this sentence does not necessarily argue for the acceptability of (65).
why (66a) is unacceptable, and why (66b) cannot have an interpretation in which the main verb of the sentence is *figure out*:

(66)  a. Joe figured that Susan wanted to take the train to New York out.
     b. Joe figured that Susan wanted to take the cat out.

For the windowing approaches like Local Association, the effect is caused by the correct head verb (*figured*) being too far from the particle (*out*). For the iterative approaches like Recency Preference, the effect is caused by the alternative head *take* which is in the way. Iterative or right-to-left locality fails as an explanation, however, because there are cases when a particle attachment is uninterpretable even if there is no possible intervening attachment point. For example, iterative models predict that a very long noun phrase without an embedded verb phrase should be interpretable, as there are no attachment points for verbal particles. However, (67a)–(67c) have no embedded verbs and yet are uninterpretable or at least quite difficult.

(67)  a. *He threw the rotten apple from the tree behind our house out.
     b. *I wrote that tedious problem set due Monday up.
     c. *I called my friend, the one from New York, up.

It is possible that a kind of windowing locality might account for the phrasal verb examples in (68)–(67). However, the window cannot be based on number of words, as Fraser (1976) argued with the examples in (68). Note that (68a) includes a four-word noun phrase between the verb and particle and is uninterpretable. But (68b)–(68d), which include interrupting noun phrases with five words, are interpretable. Thus, whatever the constraints may be on the placement of verb-particle objects, they are not statable in terms of constituent length.

(68)  a. #I called the man who left up
     b. He called all of my best friends up.
     c. Won't you total some of those larger figures up.
     d. Some charged the adding machine fire-loss off to experience.

It is possible that a kind of windowing locality might account for the phrasal verb examples in (67)–(68); note that in each of (67a)–(67c) a post-modified, phonologically heavy NP intervenes between the verb and the particle. In current work we are investigating whether a windowing locality approach together with the phonological weight criterion proposed by Zec and Inkelas (1990) (two phonological phrases) might explain the phrasal verb problem. However, significantly, even if this approach succeeds, this criterion cannot account for the different effects of the preferences in (60). The sentences in (67) are quite unacceptable; the preferences for rightmost attachment of the adverb *yesterday* in (60) are only slight preferences. Ford et al. (1982) give the following numbers for adverb attachment:
(69) Tom said Bill died yesterday.
a. Bill died yesterday. (70%)
b. Tom said it (that Bill died) yesterday. (30%)

Locality-type effects clearly play some role in a complete parsing model. But we have presented a paradox between iterative and windowing models of locality, arguing that it will be difficult to model the data with either of these approaches. In addition, we have noted that the serial parsing models which rely completely on heuristics like locality are unable to account elegantly for the delayed resolution of ambiguity needed in examples like (56).

5 SEMANTIC DISAMBIGUATION
AND A GENERAL MODEL OF INTERPRETATION

Until now we have ignored the role of semantics in disambiguation. However, a number of recent studies have shown that semantic context can reduce or eliminate the garden-path effect. Trueswell and Tanenhaus (1991), for example, show that garden path effects could be reduced by manipulating the tense of the clause. Crain and Steedman (1985) showed the effect of the semantic constraints a verb places on its arguments in examples like (70):

(70) a. #The teachers taught by the Berlitz method passed the test.
b. ?The children taught by the Berlitz method passed the test.

In current work with Srini Narayan we are investigating how our coherence model of disambiguation can be extended to deal with these semantic effects. Because we assume a sign-based or constraint-based theory of grammar, constructions are annotated with semantic information. One way of using semantic knowledge is to generalize the valence probabilities in the lexicon. Currently the probabilities refer to a solely syntactic specification of arguments. These probabilities could be extended to a function which assigns different probabilities to possible fillers of different semantic types, of the sort proposed by Resnik (1993). That is, the conceptual system of the language would be typed, and each valence slot of each predicate would be associated with a probability distribution over types. This would allow the valence probabilities for teach, for example, to distinguish teachers from children as prospective fillers. Burgess and Lund (1994) showed that computing psychological norms on similar simple thematic biases and using them to weight interpretations helped in modeling the semantic variation in garden-path sentences.

Another advantage of the probabilistic approach is that it could be extended from these kinds of 'grammatical' disambiguation examples to build a more complete theory of disambiguation. As Hirst (1986, p. 111) noted, it is impossible to disambiguate sentences like (71a,b) without nonlinguistic knowledge about "the relative aesthetics of factories and flora":
(71) a. The view from the window would be improved by the addition of a plant out there.
b. The view from the window would be destroyed by the addition of a plant out there.

A number of researchers have argued that a probabilistic model of abduction could be used to account for these sorts of nonlinguistic disambiguation, making use of probabilistic real-world knowledge. (Charniak & Goldman, 1988; Ng & Mooney, 1990; Norvig & Wilensky, 1990). Thus, by including not only valence and construction probabilities, but also conceptual and non-linguistic probabilities, a broader probabilistic theory of interpretation could model a wide-range of data on access, sentence processing, and inference.

6 PROBLEMS AND FUTURE WORK

The model as described suffers from a number of gaps and simplifying assumptions. The access algorithm assumes incorrectly that top-down and bottom-up evidence are independent. We have not faced the difficult question of morphological processing nor addressed the recently burgeoning psycholinguistic literature on morphology, or modeled overload effects like center-embedding.

In addition, the model is currently unable to account for the kind of effects that have traditionally been modeled with spreading activation. These include intra-lexical semantic effects like the priming in 'The astronomer married the star' (Reder, 1983). In addition, in an off-line experiment Gibbs, Nayak, & Cutting (1989) show effects of semantic priming in idiom processing, in which semantically decomposable idioms are processed faster than semantically nondecomposable idioms and than nonidiomatic control sentences. Our model could be fleshed out to deal with these kind of effects by adding a third kind of evidence to the access theory, some kind of semantic association evidence, in addition to top-down and bottom-up evidence.

On the other hand, these effects might also be handled on a different explanatory level of the theory, the level of activation. The model currently does not implement the time-course of construction activation. We have argued throughout the paper that access and disambiguation can be accounted for by ranking constructions by their posterior probability given the evidence; but we have not addressed implementing our model in the kind of activation or connectionist framework that is traditionally used to model low-level time-course (Elman, 1989; Feldman & Ballard, 1982; MacDonald, 1993; McClelland et al., 1989). Producing an implementation in which the time-course of activation is proportional to posterior probability (rather than frequency) remains future work; it is possible that it is at this as-yet
unaddressed level that semantic priming effects are best modeled. In addition, the work of J. Henderson (1994) in applying the connectionist paradigm of Shastri and Ajjanagadde (1993) to parsing suggests that effects like center-embedding which are commonly attributed to memory limitations could be modeled by a connectionist parser with certain memory-like limitations on variable binding. In addition, such an implementation level might allow us to incorporate the insights of competition-based approaches like that of Spivey-Knowlton and Tanenhaus (1995) or Stevenson (1993) in order to model the continuum of garden-path processing difficulty.

7 CONCLUSION

Traditional wisdom holds that a difficult problem can often be solved by divide-and-conquer methods; thus it has been argued that by dividing linguistic processing into modules for lexical, idiomatic, syntactic, and semantic processing, and orthogonally into separate models of access and disambiguation, we can eventually build a general theory of human language processing. Driven by experimental results, and resonating especially with the proposals of MacDonald (1993) and MacDonald et al. (1994), we have taken the opposite tack, proposing that a single probabilistic mechanism underlies the access and disambiguation of linguistic knowledge at every level, and demonstrating the model on psycholinguistic results at every level of linguistic structure, including lexical, idiomatic, and syntactic access and disambiguation, the interpretation of garden-path sentences, parsing preferences, and studies of gap-filling and other valence ambiguities.

Our theory makes another strong claim, regarding the use of probabilistic models in linguistic theory. Generative linguistic theory has shied away from the use of probabilistic models since Chomsky’s early arguments against Markov models of syntax. But the evidence we have presented here for the augmentation of each construction with probabilities, together with recent work which argues that probabilistic models are necessary to account for language change and learning, argues for a reanalysis of this position. Chomsky was correct in arguing against simple Markov models of syntax not because they were probabilistic, but because of their simplistic models of structure. We see probabilities not as replacements for structure, but as enrichments of structure; augmenting constructions with probabilities allows us to have the advantages of both structuralist and probabilistic models of language.

We hope this work also argues for holism at a different level, the level of academic disciplines. Building a cognitive model of parsing for a linguistic theory is necessarily an interdisciplinary enterprise. In particular, we have shown that models and metaphors from different disciplines of cognitive science can be used to solve problems in other subfields. For example, the
psycholinguistic result that human processing of language is on-line was used to solve traditional computational complexity problems in parsing. Psycholinguistic results on the strong similarities in the processing of lexical, idiomatic, and syntactic structures were used to argue for sign-based models of linguistic structure like construction grammar, cognitive grammar, or HPSG. And finally, traditional computational algorithms like beam search are used to explain psychological results.

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